## Fisheries

## Southern Zone

## Rock Lobster (Jasus edwardsif) Fishery Stock Assessment 2020/21



A. Linnane, R. McGarvey, J. Feenstra and P. Hawthorne

SARDI Publication No. F2007/000276-16 SARDI Research Report Series No. 1141

SARDI Aquatic Sciences
PO Box 120 Henley Beach SA 5022

July 2022
Fishery Assessment Report to PIRSA Fisheries and Aquaculture

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This report may be cited as:
Linnane, A., McGarvey, R., Feenstra, J. and Hawthorne, P. (2022). Southern Zone Rock Lobster (Jasus edwardsii) Fishery 2020/21. Fishery Assessment Report to PIRSA Fisheries and Aquaculture. South Australian Research and Development Institute (Aquatic Sciences), Adelaide. SARDI Publication No. F2007/000276-16. SARDI Research Report Series No. 1141. 67pp.

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29 July 2022
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## ACKNOWLEDGEMENTS

Research presented in this report was commissioned by PIRSA Fisheries and Aquaculture using funds obtained from licence fees paid by participants in the Southern Zone Rock Lobster Fishery. SARDI Aquatic and Livestock Sciences provided substantial in-kind support for the project. We thank Lachlan McLeay, Kylie Odgers and Andrew Hogg for collecting and collating the data. The report was formally reviewed by Dr Greg Ferguson, Dr Jason Earl (SARDI) and Steve Shanks (PIRSA Fisheries and Aquaculture) and approved for release by Dr Stephen Mayfield (Science Leader, Fisheries, SARDI Aquatic and Livestock Sciences).

## EXECUTIVE SUMMARY

This stock assessment determined the status of South Australia's Southern Zone Rock Lobster Fishery (SZRLF) and provides the latest estimates of the biological performance indicators (Pls) in context of the reference points (RPs) and stock status classification described in the Management Plan for the fishery (PIRSA 2020). Stock status was determined using the harvest strategy for the fishery that was developed in alignment with the National Fishery Status Reporting Framework (NFSRF) classification system that is used to determine the status of all South Australian fish stocks (Piddocke et al. 2021).

Assessment of the SZRLF relies heavily on data from the commercial fishing sector through mandatory catch and effort logbook reporting. Catch per unit effort (CPUE) of legal and undersized (pre-recruit) lobsters are the main indicators of legal and pre-recruit abundance, respectively. Fishery-independent surveys and fishery model outputs also contribute to the assessment.

The 2020 season (i.e. 2020/21) was the second consecutive year the SZRLF was impacted by overseas market disruptions. To allow for greater fishing flexibility, the 2020 season was extended from 15 September to 31 July (normally 1 October to 31 May). In 2020, the total allowable commercial catch (TACC) in the SZRLF was $1,289 \mathrm{t}$. This reflected a regular TACC of $1,246 \mathrm{t}$ plus 43 t of carry-over from the 2019 season. The total reported 2020 logbook catch was $1,275.5 \mathrm{t}$ ( $99 \%$ of TACC). Effort required to take the catch was 775,014 potlifts, the second lowest on record.

Nominal legal-sized CPUE in 2020 was 1.64 kg/potlift, reflecting a $71 \%$ increase over the last four seasons and the highest catch rate since 2004. This catch rates is now above both the long-term average and the trigger reference point (TrRP) for the fishery. Recent increases in CPUE are also reflected in fishery-independent surveys.

The pre-recruit index (PRI) shows a long-term decline since the late 1990s with the 2015 estimate the lowest on record. However, between 2015 ( 0.77 undersized/potlift) and 2019 (1.76 undersized/potlift) the PRI increased by $129 \%$, the highest since 2002. In 2020, the PRI decreased by $10 \%$ to 1.58 undersized/potlift but remained above the $\operatorname{TrRP}$ of 1.32 undersized/potlift. In the SZRLF, the time taken for pre-recruits to enter the fishable biomass is estimated to be approximately one year.

Model outputs indicate considerable increases in legal-size biomass over the last four seasons. In 2020, the estimate was approximately $4,377 \mathrm{t}$, equating to an exploitation rate of $30 \%$, the lowest on record. Despite recent increases, egg production in the fishery remains low with 2020 estimates equating to $12 \%$ of unfished levels.

In 2020, the CPUE of $1.64 \mathrm{~kg} /$ potlift was above the $\operatorname{TrRP}$ of $0.60 \mathrm{~kg} /$ potlift. As a result, the SZRLF stock is classified as "sustainable". This means that the current fishing mortality is being adequately controlled to avoid the stock becoming recruitment impaired.

Table 1 Key statistics for the SZRLF.

| Statistic | $\mathbf{2 0 2 0 / 2 1}$ | $\mathbf{2 0 1 9 / 2 0}$ |
| :--- | :---: | :---: |
| TACC | $1,289 \mathrm{t}$ | $\mathbf{1 , 2 4 5 . 7 \mathrm { t }}$ |
| Total commercial catch | $1,275.5 \mathrm{t}$ | $1,202.4 \mathrm{t}$ |
| Total effort | 775,014 potlifts | 758,029 potlifts |
| Commercial CPUE | $1.64 \mathrm{~kg} /$ potlift | $1.59 \mathrm{~kg} / \mathrm{potlift}$ |
| Pre-recruit index | 1.58 undersized/potlift | 1.76 undersized/potlift |
| Biomass estimate | $4,377 \mathrm{t}$ | $4,235 \mathrm{t}$ |
| Exploitation rate | $30 \%$ | $29 \%$ |
| Status | Sustainable | Sustainable |

Keywords: Southern Rock Lobster, Jasus edwardsii, stock assessment, harvest strategy, total allowable commercial catch.

## 1 INTRODUCTION

### 1.1 Overview

Stock assessments for the South Australian Southern Zone Rock Lobster (Jasus edwardsii) Fishery (SZRLF) have been produced annually since 1997 (Prescott et al. 1998). The current report presents information on the fishery and biology of the species and provides a current assessment of the status of the SZRLF in relation to the performance indicators provided in the Management Plan for the fishery (PIRSA 2020).

### 1.2 Description of the fishery

### 1.2.1 Access

Southern Rock Lobster is a highly valued fishery species across the States of South Australia, Victoria and Tasmania. Within South Australia, the fishery is divided into two zones: Northern and Southern, with an approximate SZRLF value of $\$ 103.5$ million in 2019/20 (BDO Econsearch 2021). The SZRLF includes all South Australian waters between the mouth of the Murray River and the Victorian border and covers an area of $22,000 \mathrm{~km}^{2}$ (Figure 1-1). It is divided into seven Marine Fishing Areas (MFAs), but the majority of fishing occurs in four MFAs ( $51,55,56$ and 58 ). There are 180 commercial licences with lobsters caught using steelframed pots (Figure 1-2) that are set overnight and hauled at first light.

### 1.2.2 Management arrangements

The SZRLF is managed by the South Australian State Government's Primary Industries and Regions South Australia (PIRSA), Fisheries and Aquaculture Division, in accordance with the legislative framework provided within the Fisheries Management (General) Regulations 2017 while specific regulations are established in the Fisheries Management (Rock Lobster Fisheries) Regulations 2017. The policy, objectives and strategies to be employed for the sustainable management of the SZRLF are described in the Management Plan for the South Australian Commercial Southern Zone Rock Lobster Fishery (PIRSA 2020). Recreational fishers are regulated under the Fisheries Management (General) Regulations 2017.

The commercial SZRLF has undergone considerable management changes over the past 50 years that have seen the fishery restructured and limited through gear restrictions, spatial and temporal closures, size limits and the implementation of a total allowable commercial catch (TACC) in 1993 (Table 1-1). The TACC is set annually and divided proportionally between licence holders owning individual transferable quota (ITQ) units. The daily catch of individual
vessels is monitored electronically via catch and disposal records and mandatory commercial logbooks. The normal fishing season extends from 1 October to 31 May of the following year. However, in 2020/21, the SZRLF was opened on 15 September 2020 and extended to 31 July 2021 to allow for greater fishing flexibility in response to overseas market issues. Full details of management arrangements for the 2020/21 season are provided in Table 1-2.

### 1.2.3 Recreational fishery

There is an important recreational fishery for lobsters in the SZRLF. Recreational fishers are permitted to use drop nets, pots or SCUBA to take lobsters during the same season as commercial fishers. All recreational lobster pots must be registered. The recreational season extends from 1 October to 31 May.


Figure 1-1 Marine Fishing Areas in the Southern and Northern Zones of the South Australian Rock Lobster Fishery.


Figure 1-2 A commercial Southern Rock Lobster fishing pot.

Table 1-1 Major management milestones for the SZRLF.

| Year | Management milestone |
| :---: | :---: |
| 1958 | Closed season for females from 1 June to 31 October and for males from 1 to 31 October |
| 1967 | Pot and boat limit introduced, no new boats to operate in the then "South-Eastern Zone" |
| 1968 | Limited entry declared, compulsory commercial catch log |
| 1978 | June, July, October closed |
| 1980 | Winter closure declared. Season from 1 October to 30 April. |
| 1984 | 15\% pot reduction |
| 1987 | Buyback of 40 licences ( 2,455 pots) |
| 1993 | April closed; TACC implemented for 1993/94 season at 1,720 t |
| 1997 | Management Plan for the fishery published (Zacharin 1997) |
| 2001 | TACC increased by 50 t to 1,770 t |
| 2003 | TACC increased by 130 t to 1,900; May opened on trial basis |
| 2005 | May trial completed. Decision to open May permanently |
| 2007 | Management Plan for the SZ fishery published (Sloan and Crosthwaite 2007) |
| 2008 | TACC reduced to 1,770 t |
| 2009 | TACC reduced to 1,400 t |
| 2010 | TACC reduced to 1,250 t. October closed to fishing |
| 2011 | New Harvest Strategy developed. October reopened |
| 2013 | Management Plan for the SZ fishery published (PIRSA 2013). One licence surrendered from fishery through marine parks voluntary commercial fisheries catch and effort reduction program |
| 2014 | TACC reduced to 1,246 to account for voluntary surrender of one licence |
| 2020 | Management Plan for the SZ fishery published (PIRSA 2020) |
| 2020 | Season opened on 15 September and extended to 31 July in response to overseas market issues |

Table 1-2 Management arrangements for the SZRLF in 2020/21.

| Management tool | Current restriction |
| :--- | :--- |
| Limited entry | 180 licences |
| Total Allowable Commercial Catch | $1,289 \mathrm{t}(1,246 \mathrm{t}+43 \mathrm{t}$ carry-over) |
| Closed season | 1 August to 14 September |
| Total number of pots | 11,882 |
| Minimum size limit | 98.5 mm CL |
| Maximum number of pots/licence | 100 pots |
| Minimum number of pots/licence | 40 pots |
| Maximum quota unit holding | Limited by pot holding (100 pots) |
| Minimum quota unit holding | Limited by minimum pot holding (40 pots) |
| Spawning females | No retention |
| Maximum vessel length | None |
| Maximum vessel power | None |
| Closed areas | Aquatic Reserves: Margaret Brock Reef, Cape Jaffa and |
|  | Rivoli Bay |
| Escape gaps | Two escape gaps or 50 mm mesh size |
| Catch and effort data | Daily logbook reported electronically |
| Catch and Disposal Records | Daily records reported electronically |
| Landing locations | 7 designated landing sites |
| Landing times | Landings permitted during core hours or outside core hours |
|  | with a prior landing report |
| Prior landing reports to PIRSA | Outside core hours, 1 hour before landing |
| Bin tags | All bins must be sealed with a lid and an approved tag prior |
|  | to lobster being unloaded from the vessel. Tags are |

### 1.3 Biology of Southern Rock Lobster

Southern Rock Lobster are distributed around southern mainland Australia, Tasmania and New Zealand. In Australia, the northern limits of distribution are Geraldton in Western Australia and Coffs Harbour in northern New South Wales but the bulk of the population is found in South Australia, Victoria, and Tasmania where they occur on algal-dominated reef habitat to depths of approximately 200 m .

Detailed reviews on the reproductive biology and life history of $J$. edwardsii are provided in Phillips (2013). In brief, J. edwardsii mate from April to July followed by a brooding period of 3-4 months over the Austral winter (June to August) (MacDiarmid 1989). Larvae hatch in early spring and pass through a brief (10-14 days) nauplius period before entering into a planktonic, leaf-like phase called a phyllosoma. These develop through a series of 11 stages over 12-23 months before metamorphosing into the puerulus stage (Booth et al. 1991; Bruce et al. 1999). Puerulus actively swim thereby aiding settlement onto suitable reef habitat (Booth et al. 1991; Phillips and McWilliam 2009).

In South Australia, the strength of westerly winds during late winter and early spring, plays an important role in inter-annual settlement variation (McGarvey and Matthews 2001; Linnane et al. 2010). After inshore settlement, early juveniles (<20 mm carapace length, CL) are solitary and normally found in isolated holes and crevices. As they develop, juvenile lobsters become increasingly communal with larger juveniles and sub-adults residing in large aggregations inside rocky dens within structurally complex reef habitat.

Based on morphological and mitochondrial DNA analysis, historical research provided little evidence of population sub-structuring across mainland Australia, Tasmania and New Zealand (Smith et al. 1980; Brasher et al. 1992; Ovenden et al. 1992). The long larval phase and widespread occurrence of larvae across the central and south Tasman Sea, in conjunction with known current flows, pointed to the likely transport of phyllosoma from south-eastern Australia to New Zealand, providing genetic mixing between the two populations (Booth et al. 1990; Bruce et al. 2007). More recent and powerful genetic techniques however have rejected the concept of panmixia and revealed significant population structure in both Tasmanian (Morgan et al. 2013) and New Zealand (Thomas 2012) stocks.

### 1.4 Research program

SARDI Aquatic Sciences maintains an on-going stock assessment and monitoring program for both the Northern and Southern Zone rock lobster fisheries of South Australia. Outputs from the program are provided to the Primary Industries and Regions South Australia (PIRSA) Fisheries and Aquaculture, through a series of annual status and stock assessment reports.

Dedicated research projects are also undertaken periodically to address to key knowledge gaps or improve ongoing stock assessments (McGarvey et al. 2014; Linnane et al. 2016).

### 1.5 Information sources for assessment

### 1.5.1 Commercial catch and effort data

All licenced commercial fishers are required to complete a daily logbook of fishing activity. This includes information such as MFA fished, species targeted, species caught, weight of legal-sized catch, number of legal-sized lobsters landed and fishing effort as potlifts. In addition to mandatory details, a number of voluntary fields may also be completed such as number of undersized individuals, lobster mortalities and levels of high-grading (total weight of lobsters returned to the water due to low market value). Records are submitted monthly to SARDI Aquatic Sciences where they are entered into the South Australian Rock Lobster (SARL) database. The catch and effort time series used in this assessment extends from 1 October 1970 to 31 July 2021.

### 1.5.2 Recreational catch and effort data

Four recreational fishing surveys have been carried out in South Australia over the past 15 years. These were primarily telephone/diary surveys in nature and were undertaken in 2000/01 (Henry and Lyle 2003), 2004/05 (Currie et al. 2006), 2007/08 (Jones 2009) and 2013/14 (Giri and Hall 2015).

### 1.5.3 Voluntary catch sampling

Since 1991, commercial fishers and researchers have collaborated in a voluntary catch sampling program. Fishers contribute by recording data from up to three pots per day while researchers generally record data from all pots during on-board observer trips. The program collects catch and effort data at finer spatial scales to that recorded in commercial logbooks in addition to supplementary data such as sex ratios, reproductive condition of females and bycatch. An important contribution from the program is lobster size data which are used to generate size frequency distributions as well as provide input data for the length-based LenMod fishery model.

### 1.5.4 Puerulus monitoring program

Rates of puerulus and post-puerulus settlement have been monitored in the SZRLF since 1991/92. This program was initiated based on the settlement-recruitment relationship observed in Western Australia where future commercial catches of Panulirus cygnus were
predicted from settlement indices using a 3-4 year time lag (Caputi et al. 1995). Though not as explicit, similar relationships are now also evident in specific regions of some J. edwardsii fisheries in both Australia and New Zealand (Gardner et al. 2001; Booth and McKenzie 2009; Linnane et al. 2013; 2014).

### 1.5.5 Fishery-Independent Monitoring Survey

It has long been recognised that fishery-dependent abundance estimates can be influenced by a range of factors associated with fishing behaviour (Thorson et al. 2017). As a result, a fishery-independent monitoring survey (FIMS) has been undertaken in the SZRLF since 2006/07. The primary aim of the FIMS is to determine legal and undersized lobster abundances, as well as size frequency distributions, that are independent of commercial fishing behaviour.

### 1.5.6 "qR" and "LenMod" stock assessment models

Two computer-based fishery stock assessment models have been developed for the South Australian Rock Lobster Fishery, referred to as "qR" and "LenMod" models. Each model provides outputs for both the Northern and Southern Zone fisheries that take into account known biological information specific to each region.

The primary data input to the qR model is catch by weight and catch by number. Outputs have been presented in stock assessment reports for the fishery since 1997 (McGarvey et al.1997; McGarvey and Matthews 2001) with a review in 2002 (Breen and McKoy 2002) concluding that the qR model was an appropriate tool for assessing rock lobster stocks. The model has been refined over time, most notably during the peer review process for publication of McGarvey and Matthews (2001) and changes to biomass definitions in 2008.

The basic structure of the second model, LenMod, was developed in the 1990s (Punt and Kennedy 1997). Variants of this length-based lobster model are now used for management and quota setting in most J. edwardsii fisheries, notably in New Zealand, Victoria and Tasmania. LenMod fits to monthly catch by number and CPUE, while conditioning on catch by weight. In addition, it also incorporates length-frequency data from voluntary catch sampling, where the lobster population is broken down into size categories of differing CL.

The primary outputs from both models are: (i) legal-sized biomass; (ii) egg production; (iii) \% unfished egg production; (iv) exploitation rate (fraction of legal-sized biomass harvested); and (v) recruitment. In addition, both models have been extensively used in bio-economic analyses and harvest strategy evaluations (McGarvey et al. 2014; 2015; 2016; 2017).

### 1.6 Harvest strategy

### 1.6.1 Management plan

A new Management Plan for the SZRLF was adopted in July 2020 (PIRSA 2020). The harvest strategy in this management plan provides a structured framework for decision-making that aims to ensure that the ecologically sustainable development objectives of the Fisheries Management Act 2007 are achieved. The aim of this harvest strategy is to improve the stock towards levels that give long-term optimum utilisation and to avoid stock over-exploitation.

### 1.6.2 Performance indicators

The Harvest Control Rule (HCR) uses multiple performance indicators to monitor the performance of the fishery (PIRSA 2020). Details of the HCR and its associated testing are provided in McGarvey et al. (2016). Broadly, the HCR aims to target a constant exploitation rate based on historical fishery performance and uses two fishery-dependent indicators.

The primary indicator is commercial CPUE (kg of legal-sized lobster/potlift) based on data from October to May, inclusive. The secondary indicator is a commercial logbook pre-recruit index (PRI; number of undersized lobsters/potlift) based on data from October to March, inclusive. Additional indicators not explicitly used to set a TACC, but which contribute to the overall assessment, include the puerulus settlement index (PSI), fishery independent monitoring survey (FIMS) outputs, length-frequency data and model outputs such as \% unfished egg production (\%UEP), exploitable biomass, exploitation rates and model-estimated recruitment.

CPUE bands, which equate to target exploitation rates, are specified in Table 1-3. To set a TACC for the upcoming season, the CPUE from the previous season is applied. A Trigger Reference Point (TrRP) of $0.60 \mathrm{~kg} /$ plotift is used, below which, exploitation rates (and corresponding TACCs) are reduced, while a Limit Reference Point (LRP) of $0.40 \mathrm{~kg} /$ potlift reflects the point at which the fishery is closed. TACCs can only be increased if the PRI is above a TrRP of 1.32 undersized/potlift. TRPs and LRPs are not applied to additional indicators.

Table 1-3 CPUE bands and associated TACCs for the SZRLF harvest control rule.

| CPUE (kg/potlift) | TACC (t) |
| :---: | :---: |
| $<0.40$ | 0 |
| $0.40-0.44$ | 337 |
| $0.45-0.49$ | 480 |
| $0.50-0.54$ | 639 |
| $0.55-0.59$ | 812 |
| $0.60-0.69$ | 897 |
| $0.70-0.79$ | 947 |
| $0.80-1.59$ | 1246 |
| $1.60-1.99$ | 1320 |
| $2.00-2.39$ | 1400 |
| $2.4+$ | 1495 |

### 1.7 Stock status classification

The status of the SZRLF was classified using the National Fishery Status Reporting Framework (NFSRF) (Flood et al. 2014) the terminology of which was recently refined and amended (Piddocke et al. 2021) (Table 1-4). It considers whether the current level of fishing pressure is adequately controlled to ensure that the stock abundance is not reduced to a point where the production of juveniles is significantly compromised. The system combines information on both the current stock size and the level of exploitation into a single classification for each stock against defined biological reference points. Each stock is then classified as 'sustainable', 'depleting', 'recovering', 'depleted', 'undefined’ or 'negligible'. PIRSA has adopted this classification system to determine the status of all key South Australian fish stocks.

The CPUE performance indicator in the current harvest strategy for the SZRLF is directly linked to a definition of stock status (Table 1-5).

Table 1-4 Stock status terminology (Piddocke et al. 2021).

| Stock status | Description | Potential implications for management of the stock |
| :---: | :---: | :---: |
| Sustainable | Biomass (or proxy) is at a level sufficient to ensure that, on average, future levels of recruitment are adequate (recruitment is not impaired) and for which fishing mortality (or proxy) is adequately controlled to avoid the stock becoming recruitment impaired (overfishing is not occurring). | Appropriate management is in place. |
| Depleting | Biomass (or proxy) is not yet depleted and recruitment is not yet impaired, but fishing mortality (or proxy) is too high (overfishing is occurring) and moving the stock in the direction of becoming recruitment impaired. | Management is needed to reduce fishing mortality and ensure that the biomass does not become depleted. |
| Recovering | Biomass (or proxy) is depleted and recruitment is impaired, but management measures are in place to promote stock recovery, and recovery is occurring. | Appropriate management is in place, and there is evidence that the biomass is recovering. |
| Depleted | Biomass (or proxy) has been reduced through catch and/or non-fishing effects, such that recruitment is impaired. Current management is not adequate to recover the stock, or adequate management measures have been put in place but have not yet resulted in measurable improvements. | Management is needed to recover this stock; if adequate management measures are already in place, more time may be required for them to take effect. |
| Undefined | Not enough information exists to determine stock status. | Data required to assess stock status are needed. |
| Negligible | Catches are so low as to be considered negligible and inadequate information exists to determine stock status. | Assessment will not be conducted unless catches and information increase. |

Table 1-5 Stock status classification for the SZRLF.

| Commercial Catch Rate (kg/potlift) | Status |
| :--- | :--- |
| $\geq 0.6$ | Sustainable |
| $<0.6$ | Depleting or Recovering |
| $\leq 0.4$ | Depleted |

## 2 METHODS

### 2.1 Commercial catch and effort statistics

Commercial logbook catch and effort data are compulsorily recorded by licenced fishers in the SZRLF. Detailed analyses of these data are provided for the period between 1 January 1970 and 31 July 2021. For ease of reference, figures and text refer to the starting year of each season (e.g. "2020" refers to the 2020/21 fishing season starting 15 September 2020).

Important commercial data such as catch (t), effort (potlifts), CPUE (kg/potlift), mean weight (kg) and PRI (number of undersized/potlift) were analysed both spatially and temporally. Spatially, data are presented by zone, MFA and in some cases, depth range. Temporally, data are presented by year and month.

In addition to the above, additional data sources recorded in the voluntary component of the logbook are also presented but at a reduced spatial or temporal scale. While these are not directly linked to setting the annual TACC, they are either deemed to contribute to the overall understanding of the fishery or have been specifically requested by stakeholder groups. These include catch rates of: (i) ovigerous (spawning) females and predation mortality as estimated through catch rates of: (ii) dead lobsters; and (iii) octopus, which are responsible for the depredation of lobsters caught in pots. The average numbers of days fished per licence holder (as a proxy for fishing effort) and estimated levels of fishery high-grading are also analysed.

### 2.2 Recreational catch and effort data

The specific details of the methodology used in the four recreational surveys considered in this assessment can be found in their respective reports (2000/01: Henry and Lyle 2003; 2004/05: Currie et al. 2006; 2007/08: Jones 2009; 2013/14: Giri and Hall 2015). A detailed description of the telephone-diary design philosophy and method is provided in Henry and Lyle (2003).

### 2.3 Voluntary catch sampling

Voluntary catch sampling datasheets are completed daily and submitted monthly to SARDI Aquatic Sciences. Fishers and observers count, measure (mm CL), and determine the sex of lobsters from all pots sampled and, for females, record the reproductive condition. In addition, all bycatch are identified and counted. The latitude and longitude of each pot sampled is recorded, thereby providing information at a finer spatial resolution than that of commercial logbooks.

### 2.4 Puerulus monitoring program

Puerulus monitoring sites in the SZRLF are located at Blackfellows Caves, Livingstones Bay, Beachport and Cape Jaffa, with the collectors set in groups of 10 or 12 at each site. The collectors are similar in design to those described by Booth and Tarring (1986) and consist of angled wooden slats that mimic natural crevice habitat. The design has remained unchanged throughout the sampling period. Sampling is undertaken monthly, whereby collector heads are detached from a base by a diver, covered with a mesh bag and hauled to the surface for counting of pueruli.

The annual puerulus settlement index (PSI) is calculated as the mean monthly settlement on all collectors combined. This index is then related to annual pre-recruit and commercial catch rate indices based on previously estimated time lags.

### 2.5 Fishery-Independent Monitoring Survey (FIMS)

The FIMS design consists of 29 transect lines, running from inshore ( $\sim 10 \mathrm{~m}$ depth) to offshore ( $\sim 120 \mathrm{~m}$ depth) fishing grounds in the SZRLF (Figure 2-1). Each transect consists of 10 pots set at predetermined locations that are independent of known fishing effort. Sampling is undertaken in September and January of each season. Lobsters are counted, sexed, measured, staged (females only) and tagged. Data are used to generate legal (number, $\mathrm{nr} / \mathrm{potlift}$ ) and undersized (nr/potlift) CPUE indices which are compared to those from fisherydependent logbook sources. FIMS abundance indices were calculated based on systematic confidence interval estimates for clustered populations (McGarvey et al. 2016 and Appendix 2).


Figure 2-1 Location of Fishery Independent Monitoring Survey (FIMS) transects in the SZRLF. MFA = Marine fishing Area.

## 2.6 " $q$ R" and "LenMod" stock assessment models

Two models are used to assess the SZRLF. The qR model is yearly and uses the three logbook time series of catch by weight, catch by number, and fishing effort as potlifts. LenMod is monthly, and integrates catch-sampling length-frequencies and FIMS CPUE, in addition to the logbook data used by the qR model. Growth in the two models differs; the qR model uses a vector of mean lengths-at-age while LenMod uses length-transition matrices. Both models estimate yearly independent recruitment. LenMod is conditioned on monthly catch-in-weight totals, while the qR model is conditioned on yearly (scaled) effort.

A number of changes were introduced to the two models in recent years. In both models, the method of computing unfished egg production (UEP) was modified by adopting 1990-2011 as the reference period for computing mean unfished recruitment. This reference period is also used in other jurisdictions (e.g. Tasmania) and therefore permits State-wide consistent reporting at the stock level under the Status of Australian Fish Stocks (SAFS) system. This reference period covers years of both higher-than-average (pre-2002) and lower-than-average (2002+) historical recruitment.

In order to remove any potential bias due to COVID-19 market impacts on fishing practices in the 2019/20 season, for the qR model, the catch rate index of abundance used is that based on reported catch and effort up to January 2020 only. Similarly, for LenMod when fitting to
catch rate and catch in number data, it does so only for October through January in 2019, and from October through May (excluding November) in 2020.

### 2.6.1 qR model

The qR model (McGarvey and Matthews 2001) fits to: (i) annual catch in weight and (ii) annual catch in number of lobsters landed. The model is effort conditioned and runs on a yearly time step. It incorporates a Baranov survival model and conditions on effort by assuming that yearly instantaneous fishing mortality rate varies in proportion to yearly reported fishing effort. The likelihood that is maximised numerically to estimate parameters is the sum of the likelihood terms for fitting to catch in weight and number. These normally distributed data provide a shared estimated parameter for the residual error as a likelihood coefficient of variation. Yearly recruitment is estimated for the start of each fishing season. Annual stock biomass is reported as an integrated average over the 12 months of each model year.

Both stock assessment models rely on commercial catch rate as a measure of relative fishable biomass. The addition of landed catches in number to the fitted logbook dataset, unavailable in most fisheries, provides important yearly information about the size of lobsters in the legal catch. Information on mean size in crustacean fisheries is normally available only from lengthfrequency samples, which can show high sample variation and are subject to additional variation in the specific locations or times during the season when length samples are taken. Catch in weight divided by catch in number gives the yearly mean weight of a landed lobster. Because reported catches in weight and number constitute a $100 \%$ sample, the quality of information obtained regarding changes in mean size from catch-log data is far more precise than that obtained from length frequencies, which typically constitute a $0.1 \%$ to $1 \%$ sample. Thus, the data informing the qR model provide relative indices of abundance as yearly catch rates (in both weight and number) and yearly mean landed weight. McGarvey et. al. (2005) demonstrated, using independent individual-based simulated data, that adding catch in number dramatically improves the accuracy and precision of stock assessment estimates in species that cannot be aged. Further details of the qR model specifications including its equations, assumptions and parameters are provided in Appendix 3.

For 2019/20, the method used to exclude qR catch rate data after January 2020 differed from that of LenMod, because the qR model is yearly. The qR model now also accounts for recreational catch. Details of these specific model updates are given in Appendix 3.

### 2.6.2 LenMod

LenMod is a length-based assessment model running on a monthly time step. Lobster population numbers are broken down and estimated in 4-mm carapace length bins. Catchability is estimated separately for each month. LenMod infers stock dynamics and abundance levels using maximum likelihood by fitting to four data sources, and conditioning on a fifth: (i) nominal monthly CPUE (in weight) to which fishable biomass is assumed to vary in direct proportion; (ii) monthly catch in number; (iii) frequency proportions by length sex bin fitted by a multinomial likelihood; (iv) fishery-independent monitoring survey (FIMS) CPUE; and ( v ) catch in weight landed from all sources (commercial and recreational). Dead lobsters in pots (predation), and lobsters dying naturally (10\% per year) are directly removed from the model population in each time step. Data sources (ii) and (iii) provide LenMod with information on the size of lobsters in the catch which, interpreted in combination with length-transition matrices, yield estimates of total mortality.

Moulting growth occurring in semi-yearly moulting times is modelled, for each sex, by lengthtransition matrices that specify the proportion of lobsters in each length class that grow into larger length classes, or remain in that length class, during each summer and autumn moulting season. These length-transition probabilities were estimated using extensive tag-recovery data mainly from the 1990s. The length-transition estimation method of McGarvey and Feenstra (2001) was applied, which infers widely flexible growth curves to be inferred by modelling the parameters predicting mean and variance of observed tag-recovery growth increments as polynomial functions of starting CL.

Growth matrices were estimated for each combination of sex and moulting season. As growth rates of female lobsters are known to slow substantially once they reach maturity, this flexible polynomial estimation method, which accommodates non-linear growth rate versus starting length, provides a more accurate estimation of female adult growth than a traditional von Bertalanffy model of mean growth increment.

Since 2018/19, the method of estimating monthly and sex-specific selectivity has been improved to allow separate length selectivity by grouped months through each season. Full details of LenMod specifications including its equations, assumptions and parameters are provided in Appendix 4.

### 2.7 Quality assurance of data

All logbook, catch sampling and fishery-independent survey data were entered and validated according to the quality assurance protocols identified for the SZRLF in the SARDI Information Systems quality assurance and data integrity report (Vainickis 2010). The data were stored in
an Oracle database, backed up daily, with access restricted to SARDI Information Systems staff. Extracts from the database were provided to SARDI rock lobster researchers on request. All puerulus data were entered into Excel spreadsheets and stored on a SARDI network drive

## 3 RESULTS

### 3.1 Commercial catch and effort statistics

The 2020 season (i.e. 2020/21) was the second consecutive year the SZRLF was impacted by overseas market disruptions. This was particularly evident in November of 2020 but should be taken into consideration when interpreting fishery trends throughout the entire 2020/21 season.

### 3.1.1 Zone

In 2020, the TACC in the SZRLF was $1,289 \mathrm{t}$. This reflected a regular TACC of $1,246 \mathrm{t}$ plus 43 t of carry-over from the 2019 season. The total reported logbook catch (September-July) was $1,275.5 \mathrm{t}$ ( $99 \%$ of TACC) (Figure 3-1a; Table 5-1). Current catch levels are low in a historical context and reflect declines in catch and subsequent TACC reductions from 2007 to 2009 (Table 3-1). During this timeframe, the TACC was reduced by $34 \%$ from $1,900 \mathrm{t}$ to $1,250 \mathrm{t}$ with a further reduction to $1,245.7 \mathrm{t}$ in 2014 due to the removal of one licence as part of the marine parks voluntary commercial fisheries catch and effort reduction program. Catches have been stable since 2014, reflecting the constant TACC over this period.

Effort required to take the 1,275.5 t catch was 775,014 potlifts, an increase of $2 \%$ from 2019 ( 758,029 potlifts) (Figure 3-1a; Table 5-1). Since 2009 ( $2,049,961$ potlifts), effort has generally declined in the fishery, with the 2020 estimate being the second lowest on record.

In 2020, the nominal legal-sized CPUE (October-May) was $1.64 \mathrm{~kg} / \mathrm{potlift}$, reflecting a $71 \%$ increase over the last four seasons (from $0.96 \mathrm{~kg} /$ potlift in 2016) and the highest catch rate since 2004 (Figure 3-1b; Table 5-1). Between 2010 and 2016, catch rates remained relatively stable at approximately $1 \mathrm{~kg} /$ potlift. The 2020 estimate represents the fifth time since 2007 that CPUE has been above the long-term average ( $1.06 \mathrm{~kg} / \mathrm{potlift}$ ). CPUE in 2020 also remains above the trigger reference point ( TrRP ) of $0.60 \mathrm{~kg} /$ potlift.

Overall, the zonal estimate of the logbook-based PRI (October-March) shows a long-term decline between 1999 and 2015 (Figure 3-1c). However, between 2015 ( 0.77 undersized/potlift) and 2019 ( 1.76 undersized/potlift) the PRI increased by 129\%, the highest since 2002. In 2020, the PRI decreased by $10 \%$ to 1.58 undersized/potlift but remained above the TrRP of 1.32 undersized/potlift. In the SZRLF, the time taken for pre-recruits to enter the fishable biomass is estimated to be approximately one year.

Legal-sized mean weight has remained relatively stable over time ranging between 0.7 and 0.9 kg (Figure 3-1d). However, over the last three seasons the mean weight increased and in 2020, was 0.96 kg , the highest on record. Variations in mean weight generally reflect long-
term patterns of recruitment, with low mean weights resulting from influxes of small lobsters into the fishable biomass and high mean weights representing several consecutive years of reduced recruitment.


Figure 3-1 Fishery dependent outputs for the SZRLF. (a) Catch and effort including total allowable commercial catch (TACC) limit; (b) catch per unit effort (CPUE) including long-term average (dashed line); (c) pre-recruit index (PRI) including trigger reference point (dashed line); and (d) mean weight.

Table 3-1 Chronology of TACC versus landed catch in the SZRLF. *Includes 43 t carry-over from 2019/20.

| Season | TACC $(t)$ | Landed catch $(t)$ | Shortfall $(t)$ | \% TACC taken |
| :---: | :---: | :---: | :---: | :---: |
| 1993 | 1720 | 1668.6 | 51.4 | 97 |
| 1994 | 1720 | 1721.5 | -1.5 | 100 |
| 1995 | 1720 | 1683.6 | 36.4 | 98 |
| 1996 | 1720 | 1639.7 | 80.3 | 95 |
| 1997 | 1720 | 1680.0 | 40.0 | 98 |
| 1998 | 1720 | 1713.1 | 6.9 | 100 |
| 1999 | 1720 | 1717.3 | 2.7 | 100 |
| 2000 | 1720 | 1716.3 | 3.7 | 100 |
| 2001 | 1770 | 1717.5 | 52.5 | 97 |
| 2002 | 1770 | 1765.9 | 4.1 | 100 |
| 2003 | 1900 | 1895.9 | 4.1 | 100 |
| 2004 | 1900 | 1896.6 | 3.4 | 100 |
| 2005 | 1900 | 1888.7 | 11.3 | 99 |
| 2006 | 1900 | 1893.9 | 6.1 | 100 |
| 2007 | 1900 | 1849.6 | 50.4 | 97 |
| 2008 | 1770 | 1407.3 | 362.7 | 80 |
| 2009 | 1400 | 1243.3 | 156.7 | 89 |
| 2010 | 1250 | 1244.2 | 5.8 | 100 |
| 2011 | 1250 | 1242.1 | 7.9 | 99 |
| 2012 | 1250 | 1234.4 | 15.6 | 99 |
| 2013 | 1250 | 1246.7 | 3.3 | 100 |
| 2014 | 1245.7 | 1244.4 | 1.3 | 100 |
| 2015 | 1245.7 | 1244.4 | 1.3 | 100 |
| 2016 | 1245.7 | 1237.7 | 8.0 | 99 |
| 2017 | 1245.7 | 1245.7 | 0 | 100 |
| 2018 | 1245.7 | 1245.2 | 0.5 | 100 |
| 2019 | 1245.7 | 1202.4 | 43.3 | 96 |
| 2020 | $1289^{*}$ |  | 1275.5 | 13.5 |

### 3.1.2 Within-season trends

Within-season commercial catch trends presented here are based on logbook data from 2018 to 2020. Results from earlier seasons are accessible through previously published stock assessment reports (http://pir.sa.gov.au/research/publications/research_reports). In general, within-season trends in catch, effort, CPUE, PRI and mean weight within the SZRLF are consistent through time although the impact of overseas market disruption closures on catch, effort and CPUE trends are evident in both February of the 2019/20 season and November of the 2020/21 season (Figure 3-2a-d). The highest catches are taken during spring/summer from October to January, before declining thereafter.

In 2020, in response to overseas market issues, the fishery opened on 15 September with 183 t landed during this month. The market closure occurred in November of 2020 and consequently, the catch in November decreased to 39 t , where normally up to 250 t are landed (Figure 3-2a). Catch subsequently increased as market conditions improved between December and March ( 696 t taken over this period). In 2020, the highest catch was taken in October 2020 ( 301 t ), and the lowest catch in July 2021 ( <1 t). Within-season effort levels are largely consistent with those of catch (Figure 3-2a). In 2020, effort was highest in October (204,633 potlifts) and lowest in July (<1000 potlifts).

Legal-sized CPUE within the fishery is generally highest at the start of the season in spring/summer before decreasing thereafter (Figure 3-2b). In 2020, monthly catch rates were consistently highest (>1.7 kg/potlift) from December to March before decreasing thereafter. In 2020, CPUE was highest in February ( $1.82 \mathrm{~kg} / \mathrm{potlift}$ ) and lowest in June ( $0.82 \mathrm{~kg} / \mathrm{potlift}$ ).

In 2020, with the exception of November and December, estimates of PRI were consistently lower across all months compared to 2019 (Figure 3-2c). The PRI was highest in November ( 1.99 undersized/potlift) and lowest in May ( 0.53 undersized/potlift).

Monthly legal-sized mean weight generally increases as the season progresses with trends broadly similar over the last three seasons, except for May 2018, although care should be taken when interpreting this estimate due to low catch levels (Figure 3-2d). In 2020, mean weight was lowest in November ( 0.87 kg ) and highest in May ( 1.19 kg ).


Figure 3-2 Within-season fishery dependent trends in the SZRLF. (a) Catch and effort; (b) catch per unit effort (CPUE); (c) pre-recruit index (PRI); and (d) mean weight.

### 3.1.3 Spatial trends

### 3.1.3.1 Marine Fishing Areas (MFAs)

Over $95 \%$ of the catch in the SZRLF is taken from MFAs 55, 56 and 58 (Figure 1-1). Historically, MFA 51 was a more important area, but its contribution has decreased in recent seasons. This partially reflects the fact that lobsters harvested from MFA 51 are generally larger in size and have a lower market value given the preference for smaller individuals by overseas markets. In 2020, the catches in MFAs 51, 55,56 and 58 were $35 \mathrm{t}, 472 \mathrm{t}, 411 \mathrm{t}$ and 331 t, respectively (Figure 3-3a; Table 5-2).

In 2020, effort increased in MFAs 51, 55 and 56 but decreased in MFA 58. Effort estimates in 2020 in MFAs 51, 55,56 and 58 were $22,768,274,676,258,259$, and 205,334 potlifts, respectively (Figure 3-3a; Table 5-2). These estimates reflect considerable decreases in effort across all areas, particularly over the last three to four seasons.

Trends in CPUE are temporally consistent across MFAs (Figure 3-3b;Table 5-2). Following considerable declines between 2002 and 2009, catch rates remained relatively stable between 2010 and 2016. Over the last three seasons, however, catch rates have increased across all major MFAs. In 2020, CPUE decreased in MFAs 51, 55 and 56 but increased in MFA 58. The 2020 estimates in MFAs 51, 55, 56 and 58 were $1.53,1.70,1.60$ and $1.61 \mathrm{~kg} /$ potlift respectively.

Spatial estimates of the logbook-based PRI indicate that the number of undersized/potlift is consistently lower in the northern regions of the SZRLF (i.e. MFAs 51 and 55) compared to the southern regions (i.e. MFAs 56 and 58) (Figure 3-3c). Estimates have been increasing in most MFAs over the last four seasons but with marginal decreases in 2020 in all MFAs except for MFA 51. In 2020, the estimates were $0.30,0.45,1.87$ and 3.02 undersized/potlift, in MFAs $51,55,56$ and 58 respectively.

Rock lobster legal-sized mean weight decreases with increasing latitude from the mouth of the Murray River (MFA 51) to the Victoria/South Australia border (MFA 58) (Figure 3-3d). It is most variable in MFA 51 but generally consistent across other MFAs. In 2020, mean weight increased in all areas with estimates in MFAs 51, 55,56 and 58 , being 1.33, 1.21, 0.88 and 0.77 kg , respectively.


Figure 3-3 Spatial fishery dependent trends in the SZRLF. (a) Catch and effort; (b) catch per unit effort (CPUE); (c) pre-recruit index (PRI); and (d) mean weight.

### 3.1.3.2 Depth

In order to assess spatial trends by depth, logbook-derived catch data from four depth range categories of $0-30,31-60,61-90$ and $>90 \mathrm{~m}$ were analysed. Over the last sixteen fishing seasons, over $80 \%$ of the annual catch has been taken from depths of $<60 \mathrm{~m}$ (Figure 3-4). In $2020,96 \%$ of the total catch came from $\leq 60 \mathrm{~m}$ depth with $36 \%$ coming from $0-30 \mathrm{~m}$ and $60 \%$ from 31-60 m. These trends were also reflected in each of the major MFAs (Figure 3-4).

Despite reflecting the majority of annual catches, CPUE in depths of $0-30 \mathrm{~m}$ and $31-60 \mathrm{~m}$ is consistently lower than offshore areas of 61-90 m and $>90 \mathrm{~m}$ (Figure 3-5). Trends largely reflected those at the zonal level with considerable decreases in all depth ranges from 2002 to 2009 before generally increasing over the last decade. In 2020, estimates were 1.64, 1.62, and $2.17 \mathrm{~kg} /$ potlift in $0-30,31-60$, and $61-90 \mathrm{~m}$, respectively.


Figure 3-4 Percentage of catch taken from four depth ranges in the SZRLF by zone (top) and across the primary MFAs (bottom) from 2003 to 2020.


Figure 3-5 CPUE by depth in the SZRLF from 1970 to 2020.

### 3.1.4 Additional indices

### 3.1.4.1 Ovigerous (spawning) females

In 2020, the catch rate of ovigerous (spawning) lobsters (September to July) was 0.62 spawners/potlift, the highest estimate on record (Figure 3-6a). In line with overall declines in legal-sized lobster catch rates (Figure 3-1b), the CPUE of spawners decreased from 2002 to a historical low of 0.05 spawners/potlift in 2010. Since then, the index has been variable, with notable increases over the last three seasons. It is important to note that as October was closed for the 2010 season, the CPUE for spawning lobsters in that season is likely to be underestimated since October is commonly the highest catch month for ovigerous individuals.

### 3.1.4.2 Predation mortality

The maori octopus (Pinnoctopus cordiformis) is the primary predator of Southern Rock Lobster within commercial fishing pots (Brock and Ward 2004). As a result, both the catch rate of octopus and dead lobsters appear correlated (Figure 3-6b; $\mathrm{R}^{2}=0.63$ ). The number of dead lobsters/potlift has been variable through time ranging from 0.09 dead/potlift (in 2009) to 0.27 dead/potlift (in 2004) (Figure 3-6b). In 2020, the estimate was 0.22 dead/potlift. The highest octopus catch rate was observed in 2000 at 0.05 octopus/potlift, with the lowest in 2017 at 0.008 octopus/potlift (Figure 3-6b). In 2020, the estimate was 0.02 octopus/potlift, which is one of the lowest on record for the fishery.

### 3.1.4.3 Average days fished

In 2020, the average number of days fished per licence in the SZRLF was 60 ( $\pm 18$ SD) days, the lowest estimate on record (Figure 3-6c). This index is a proxy for overall fishing effort and largely reflects trends in annual potlifts within the fishery (Figure 3-1a). From 2004 to 2009, the average number of days fished increased by $86 \%$ from 94 to 175 , the highest on record, despite reductions to the TACC from $1,900 \mathrm{t}$ to $1,400 \mathrm{t}$ over the same period. In 2010, the TACC was reduced to $1,250 \mathrm{t}$ and the average numbers of days fished decreased to 114 days, the lowest since 2005 (105 days). In 2013, the TACC was further reduced to 1245.7 t under the marine parks voluntary commercial fisheries catch and effort reduction program. The TACC has remained at this level over the last eight seasons.

### 3.1.4.4 High-grading

In 2020, the estimate of high-grading (i.e. weight lobsters returned to the water due to low market value) in the SZRLF was 63 t (Figure 3-6d). From 2003 to 2006, based on voluntary catch returns, the weight of lobsters high-graded exceeded 100 t annually. However, 2019 was the first season since 2008, that the estimates have exceeded 30 t . The decrease between 2003 and 2008 is likely to reflect overall declines in legal-sized catch rate across the fishery over this period (Figure 3-1b). It should be highlighted that overall reported values in logbooks are likely to be conservative, since high-grade estimates are recorded on a voluntary basis.


Figure 3-6 Additional fishery dependent indices in the SZRLF. (a) Catch rate of spawning lobsters; (b) predation mortality; (c) average number of days fished (with S.D.); and (d) levels of high-grading.

### 3.2 Recreational catch and effort

The most recent report on recreational rock lobster fishers was undertaken during the 2013/14 South Australian Recreational Fishing Survey (Giri and Hall 2015). An estimated 102,931 $( \pm 58,763)$ lobsters were caught by South Australian residents with $62,346( \pm 39,085)$ of these harvested and $40,585( \pm 25,202)$ released representing a release rate of $39.4 \%$. In total, the harvested catch equated to approximately 75 t of which two-thirds were caught in the SZRLF. Pots and nets accounted for $83 \%$ of all lobsters caught with dive fishing being the other major capture method.

These results can be compared with 106,483 lobsters caught in 2007/08 with 47,875 harvested (equating to 60 t with approximately 55 t caught in the SZRLF) and 58,608 released representing a release rate of $55 \%$ (Jones 2009). A 2021-22 recreational fishing survey will provide State-wide estimates of rock lobster catch from 1 March 2021 to 28 February 2022, with results due to be released by the end of 2022. Recreational catches are accounted for within LenMod fishery outputs.

### 3.3 Voluntary catch sampling

Since 1991, up to 26,000 lobsters have been measured annually in the SZRLF as part of the voluntary catch sampling program. The number measured is proportional to the level of participation in the program with data presented as number of lobsters/100 potlifts. In this report, annual length frequency data are presented from 2011-2020. Earlier annual length frequency distributions are available in published stock assessment reports (http://pir.sa.gov.au/research/publications/research_reports).

Male lobsters, which generally grow faster and reach larger sizes than females, range between 70 and 200 mm carapace length (CL). In contrast, few females are larger than 150 mm CL. In 2020, a total of 14,462 lobsters were sampled. Length-frequency data obtained through the voluntary catch sampling program over the last two seasons (Figure 3-7) support recent trends in pre-recruit indices (Figure 3-1c). Notably, the percentage of lobsters measured below the minimum legal size (MLS) of 98.5 mm CL increased from $23 \%$ in 2015 to $43 \%$ in 2019, reflecting the increase in undersized catch rate over the same period (Figure 3-1c). The decrease to $29 \%$ between 2019 and 2020 also reflects the decline in PRI over the two seasons.


Figure 3-7 Length frequency distributions of male and female lobsters combined in the SZRLF from 2011 to 2020 (red line indicates MLS at 98.5 mm CL).

### 3.4 Puerulus monitoring program

Puerulus settlement indices (PSIs) in the SZRLF have been highly variable over time (Figure $3-8$ ). In 2020, the estimate was 2.23 puerulus/collector, which is above both the median and long-term average for the fishery. Previous research has indicated that the period between settlement and recruitment to legal size in the SZRLF is approximately five years with undersized numbers correlated after four years (Linnane et al. 2014).

Based on this relationship, PSIs were correlated against logbook PRI and CPUE data lagged by four and five years respectively, using PSI data from 1991 to 2016 (Figure 3-9). CPUE and PRI data were closely correlated over the entire time series with a one-year lag ( $R^{2}=0.73$ ) confirming that most pre-recruits enter the fishery one year later. PSIs from 2002 to 2016 were poorly correlated ( $\mathrm{R}^{2}=0.34$ ) with subsequent PRIs from 2006 to 2020 with no obvious relationship between PSI and CPUE over any of the time series.

Three of the highest PSIs on record were observed from 2005 to 2007 which were reflected in subsequent increases in PRIs in 2009 and 2010 (Figure 3-1c) and increases in catch rate as these recruits entered into the fishery in 2010 and 2011 (Figure 3-1b). However, these relationships are not consistent throughout the time series. For example, increases in CPUE in 2017 and 2018 came from historically low settlements in 2012 and 2013. Recent trends over the last seven seasons indicate that recruitment to the fishery will be close to the longterm average in the short-to-medium term.


Figure 3-8 Puerulus settlement indices (mean $\pm$ SE) in the SZRLF from 1991 to 2020. Dashed and solid black lines represent long-term (1991-2020) average and median estimates respectively.


Figure 3-9 Correlations between SZRLF puerulus settlement lagged by four years with logbook PRI.

### 3.5 Fishery-Independent Monitoring Survey (FIMS)

The latest FIMS for the SZRLF 2021 season was completed in February 2022. Between 2006 and 2015, with the exception of 2015 for legal-size lobsters, the survey catch rates of both legal-size (Figure 3-10) and undersized (Figure 3-11) lobsters generally decreased. From 2016 to 2021 (except for 2019), the catch rate of legal-size lobsters increased from 0.46 to 1.26 lobsters/potlift, with the 2021 estimate the highest on record. With the exception of 2018, the catch rate of undersized lobsters also increased from 2015 ( 0.17 undersized/potlift) to 2020 ( 0.48 undersized/potlift). In 2021, the estimate was 0.50 undersized/potlift).

Trends in fishery-independent indices were compared against those from the commercial fishery. For legal-sized lobsters, both commercial and survey indices decreased from 2006 to 2009. However, after this period the trends for these indices diverged. In 2010, commercial catch rate increased and remained relatively stable at approximately 1.2-1.4 lobsters/potlift. In contrast, between 2008 and 2016, legal-sized survey CPUE remained at historically low levels ranging from 0.43 to 0.74 lobsters/potlift (Figure 3-10). Both indices have generally increased over the last 4 to 5 seasons.

Trends in survey and commercial undersized CPUE were similar between 2006 and 2009 (Figure 3-11). Since then, the rate of decline in abundances of undersized lobsters in surveys was greater than that observed in commercial catch rate data. Both survey and commercial data showed general increases in undersized catch rates between 2015 and 2020.


Figure 3-10 Comparison of legal size catch rates (nr/potlift) from commercial logbook data and fishery independent monitoring surveys.


Figure 3-11 Comparison of undersized catch rates (nr/potlift) from commercial logbook data and fishery independent monitoring surveys.

## 3.6 "qR" and "LenMod" stock assessment models

### 3.6.1 Model fits

Both the $q R$ and LenMod fishery model show good fits to the available data (Appendix 5). The qR model fitted closely to logbook totals of yearly catch in number (Figure 5-1) and catch in weight (Figure 5-2). For LenMod, model estimates of monthly catch in number and catch rate fitted closely to the reported monthly logbook catch in number ( Cn ) (Figure 5-3) and catch rate (Figure 5-4). In addition, model estimates for both males and females fitted well to lengthfrequency data from voluntary catch sampling as shown in monthly fits from the 2020 season (Figure 5-5).

### 3.6.2 Model outputs

The SZRLF qR and LenMod models show close agreement in estimated trends for indicators of performance and status. From 2002 to 2009, estimates of legal-sized biomass decreased by $62 \%$, from approximately $5,000 \mathrm{t}$ to $1,900 \mathrm{t}$ (Figure $3-12 \mathrm{a}$ ). Since then, biomass has increased, particularly over the past four seasons, and in 2020, the estimate was approximately $4,377 \mathrm{t}$.

Coincident with declines in lobster biomass, egg production estimates decreased by $52 \%$ from approximately 650 billion in 2003 to 310 billion in 2009 (Figure 3-12b). Over the last eleven seasons, egg production has trended upward, and in 2020 was estimated at approximately 547 billion. However, despite recent increases, overall egg production estimates are low with current estimates equating to 12\% of unfished levels (Figure 3-12c).

Exploitation rate increased from approximately 37\% in 2002 to 73\% in 2009 (Figure 3-12d) in response to decreasing biomass over the same period (Figure 3-12a). Exploitation rate decreased considerably in 2010 and has continued to gradually decline since. In 2020, after four years of large reductions in exploitation rate reflecting corresponding increases in biomass, the estimate was approximately $30 \%$, which is the lowest on record.

Model outputs indicate that recruitment to the fishery declined from approximately 4 million individuals in 1999 to 1 million in 2008, a decrease of $75 \%$ (Figure 3-12e). Since then, the estimate has been variable and in 2020 ranged from 1.7-2.7 million individuals. Temporal trends in recruitment estimated by both models are correlated $\left(R^{2}=0.82\right)$ with PRI from logbook data from 1995-2020 (Figure 3-12e).


Figure 3-12 Fishery model outputs for the SZRLF. (a) Legal-size biomass; (b) Egg production; (c) \% of unfished egg production; (d) Exploitation rate; and (e) Recruitment

## 4 DISCUSSION

### 4.1 Information sources used for assessment

Assessment of the SZRLF resource relies heavily on commercial fishery-dependent data collected from several long-term monitoring programs. It places particular emphasis on assessing catch rate trends of both legal and undersized lobsters. These trends are supported by fishery-independent surveys and outputs from both the qR and LenMod fishery models.

Current catch rates are not standardised. Linnane et al. (2018) presented standardised estimates which were reviewed by the South Australian rock lobster Harvest Strategy Working Group (HSWG) which noted the close agreement between nominal and standardised timeseries. The HSWG recommended that periodic catch rate standardisation should be continued, but that nominal catch rate could remain as the primary indicator of lobster abundance.

### 4.2 Stock Status

The 2020 season reflected the second consecutive year the SZRLF was impacted by overseas market issues due to a combination of international trade disputes and COVID. The impact can be observed in within-season trends e.g. catch in February of 2020 was reduced to 6 t when normally up to 100 t of lobster are landed, while in November 2020 just 39 t was taken when normally up to $250 t$ is caught. These impacts can affect important indicators such as catch rate, at least on a monthly level. For example, the low CPUE estimate observed in February 2020 is clearly a reflection of known market influences, rather than a reduction in lobster abundance. Despite these impacts, there are clear signals to indicate that the status of the SZRLF has improved in recent seasons.

Up to 2020, the SZRLF TACC had been fully taken for eleven consecutive seasons. The current catch level of $1,275.5 \mathrm{t}$ is low in a historical context with effort required to take the TACC having generally decreased since 2010. Reduced levels of catch have resulted in a considerable improvement in SZRLF performance, particularly since 2015. Legal-size catch rates have increased by 71 \% over the last four seasons and are now the highest since 2004. CPUE is now above both the long-term average and TRP with increases observed across broad spatial and temporal scales. Similar increases in legal-size catch rate have also been observed over recent seasons in fishery-independent surveys.

As well as reduced catch levels, recent increases in catch rate are likely driven by improved recruitment to the fishable biomass, particularly over the last four seasons. After a long-term decline from 1999 to 2015, undersized abundances have increased and are now the highest since 2002. Given that the period between PRI and recruitment to the fishery is approximately one year, this has translated to increases in legal-size CPUE. Recent recruitment increases are supported by independent model estimates of recruitment, undersize length frequency distributions and fishery-independent PRI estimates.

In the SZRLF, based on tag-recapture studies, the period between puerulus settlement and the PRI is $\sim 4$ years, with recruitment into the fishery occurring one year later (i.e. 5 years after settlement) (McGarvey et al. 1999; Linnane et al. 2013; 2014). While correlations between the PRI and CPUE are strong using a one-year lag, consistent correlations between settlement and recruitment are yet to emerge. For example, high levels of settlement observed in 2005, 2006 and 2009 reflected the increase in PRI in 2009, 2010 and 2013 which, combined with reduced catch levels, is likely to have contributed to elevated commercial CPUE over recent seasons. However, these relationships are not consistent over time as highlighted by the fact that the recent increases in CPUE in 2017 and 2018 came from below average settlements in 2012 and 2013.

Outputs from the qR and LenMod fishery models agree closely in relation to the current stock status. Following a considerable decline between 2002 and 2009, legal-sized biomass has gradually increased over the last eleven seasons. Given that catch has remained stable over this period, exploitation rate has decreased, with the 2020 estimate the lowest on record. Egg production levels have also increased in recent seasons but remain low in 2020 at $12 \%$ of \%UEP.

In 2019, a new Management Plan for the SZRLF was formally adopted (PIRSA 2020). To address low levels of \%UEP, the harvest strategy focuses on increasing egg production towards a stock improvement target of $20 \%$ by 2035. Importantly, the primary performance indicator (legal-size CPUE) in this harvest strategy is now linked to a definition of stock status (Table 1-5). In 2020, the CPUE was $1.64 \mathrm{~kg} /$ potlift, which is above the $\operatorname{TrRP}$ of $0.60 \mathrm{~kg} /$ potlift. As a result, the SZRLF stock is classified as "sustainable". This means that the current fishing mortality is being adequately controlled to avoid the stock becoming recruitment impaired.

### 4.3 Assessment Uncertainties

This assessment is reliant on fishery-dependent data as an indicator of stock abundance. However, it is widely acknowledged that catch rate estimates, based on fishery-dependent data, can be influenced by factors such as gear selectivity, changes in fishing patterns, fleet
efficiency or fleet dynamics over time (Maunder et al. 2006). Nevertheless, two lines of evidence suggest that the catch rate trends detailed in this report are robust indicators of overall lobster abundance. Firstly, trends are highly consistent across large spatial scales. For example, across the four major MFAs of the fishery, catch rate simultaneously increased from the mid-1990s to early 2000s, declined to historical lows from 2002 to 2009, before gradually recovering over the next eight seasons. Similar trends were also observed across a range of depth categories within MFAs. These fishery-wide trends suggest recruitment and subsequent survival in the SZRLF occur consistently across large spatial scales, and that these trends are well reflected in the broad seasonal and spatial coverage (>1 million potlifts annually) used to compute catch rate.

Secondly, a previous stock assessment report (Linnane et al. 2018) highlighted that when nominal catch rate was standardised for factors such as year, month, depth, MFA, mean weight, licence and consumer price index (CPI), the nominal and standardised CPUE time series were closely aligned. While no meaningful difference, and therefore no improvement was observed, the standardisation did not include two factors thought to be important in other lobster fisheries. Specifically, standard catch logs in South Australian lobster fisheries do not record the "vessel" or "skipper". In the Victorian rock lobster fishery, "vessel" and "skipper" were identified as the two most important factors in legal-size lobster catch rate standardisation (Feenstra et al. 2019).

### 4.4 Future Work

The uncertainties with overseas markets have made research planning difficult over the past two seasons. This is because many research projects require industry collaboration in terms of vessel time/support which cannot be guaranteed in the current market environment.

There is a need to investigate changes in growth rates within the fishery over time. Most growth information in the models is based on the large-scale tag/recapture program undertaken from 1993 to 1996 (Linnane et al. 2005). Quantifying a change in growth rate will require relaunching a substantial tag-recovery program. This has been identified as a high research priority given the importance of accurate growth transition matrices and mean weight-at-age in the estimation of absolute biomass within the two fishery assessment models.

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## 5 APPENDICES

## Appendix 1. SZRLF Catch, Effort and CPUE data

Table 5-1 Catch, Effort and CPUE (commercial and fishery independent survey) for the SZRLF from 1970 to 2020 by zone.

|  |  |  | Commercial | Survey |
| :---: | :---: | :---: | :---: | :---: |
| Season | Catch (t) | Effort (000's potlifts) | CPUE (kg/potlift) | CPUE ( $\mathrm{Nr} /$ potlift) |
| 1970 | 1544 | 1838 | 0.84 |  |
| 1971 | 1604 | 1987 | 0.81 |  |
| 1972 | 1949 | 2138 | 0.91 |  |
| 1973 | 1738 | 2135 | 0.81 |  |
| 1974 | 1212 | 1727 | 0.70 |  |
| 1975 | 1621 | 2006 | 0.81 |  |
| 1976 | 1374 | 1882 | 0.73 |  |
| 1977 | 1300 | 1875 | 0.69 |  |
| 1978 | 1309 | 1801 | 0.73 |  |
| 1979 | 1534 | 1793 | 0.86 |  |
| 1980 | 2126 | 2029 | 1.05 |  |
| 1981 | 2047 | 2122 | 0.96 |  |
| 1982 | 1844 | 2162 | 0.85 |  |
| 1983 | 1734 | 2255 | 0.77 |  |
| 1984 | 1537 | 1984 | 0.77 |  |
| 1985 | 1547 | 2020 | 0.77 |  |
| 1986 | 1458 | 1909 | 0.76 |  |
| 1987 | 1657 | 2130 | 0.78 |  |
| 1988 | 1407 | 1886 | 0.75 |  |
| 1989 | 1528 | 1798 | 0.85 |  |
| 1990 | 1563 | 1907 | 0.82 |  |
| 1991 | 1940 | 2050 | 0.95 |  |
| 1992 | 1754 | 1759 | 1.00 |  |
| 1993 | 1669 | 1642 | 1.02 |  |
| 1994 | 1721 | 1511 | 1.14 |  |
| 1995 | 1684 | 1591 | 1.06 |  |
| 1996 | 1640 | 1755 | 0.93 |  |
| 1997 | 1680 | 1758 | 0.96 |  |
| 1998 | 1713 | 1537 | 1.11 |  |
| 1999 | 1717 | 1162 | 1.48 |  |
| 2000 | 1716 | 1039 | 1.65 |  |
| 2001 | 1717 | 910 | 1.89 |  |
| 2002 | 1766 | 854 | 2.07 |  |
| 2003 | 1896 | 1042 | 1.82 |  |
| 2004 | 1897 | 1052 | 1.80 |  |
| 2005 | 1889 | 1183 | 1.60 |  |
| 2006 | 1894 | 1354 | 1.40 | 1.24 |
| 2007 | 1850 | 1661 | 1.11 | 1.00 |
| 2008 | 1407 | 1916 | 0.73 | 0.54 |
| 2009 | 1243 | 2050 | 0.61 | 0.51 |
| 2010 | 1244 | 1322 | 0.94 | 0.65 |
| 2011 | 1242 | 1285 | 0.97 | 0.54 |
| 2012 | 1234 | 1419 | 0.87 | 0.49 |
| 2013 | 1247 | 1253 | 1.00 | 0.43 |
| 2014 | 1244 | 1207 | 1.03 | 0.46 |
| 2015 | 1244 | 1220 | 1.02 | 0.74 |
| 2016 | 1238 | 1296 | 0.96 | 0.46 |
| 2017 | 1246 | 1022 | 1.22 | 0.75 |
| 2018 | 1245 | 841 | 1.48 | 0.96 |
| 2019 | 1202 | 758 | 1.59 | 1.19 |
| 2020 | 1276 | 775 | 1.64 | 1.26 |

Table 5-2 Catch, Effort and CPUE for the SZRLF from 1970 to 2020 by MFA.

|  | MFA 51 |  |  | MFA 55 |  |  | MFA 56 |  |  | MFA 58 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Season | Catch (t) | Effort (000's potifits) | CPUE (kg/potilift) | Catch (t) | Effort (000's potilifs) | CPUE (kg/potilift) | Catch (t) | Effort (000's potilifs) | CPUE (kg/potift) | Catch (t) | Effort (000's potlifts) | CPUE (kg/potilif) |
| 1970 | 215 | 219 | 0.91 | 443 | 489 | 0.91 | 486 | 616 | 0.79 | 377 | 499 | 0.82 |
| 1971 | 267 | 285 | 0.92 | 455 | 510 | 0.92 | 494 | 663 | 0.78 | 374 | 529 | 0.69 |
| 1972 | 298 | 305 | 0.98 | 588 | 600 | 0.97 | 632 | 672 | 0.97 | 403 | 543 | 0.72 |
| 1973 | 224 | 280 | 0.80 | 488 | 570 | 0.85 | 582 | 676 | 0.88 | 418 | 591 | 0.71 |
| 1974 | 166 | 221 | 0.78 | 331 | 456 | 0.72 | 405 | 567 | 0.72 | 309 | 481 | 0.63 |
| 1975 | 220 | 244 | 0.89 | 418 | 517 | 0.81 | 571 | 691 | 0.81 | 403 | 543 | 0.74 |
| 1976 | 184 | 226 | 0.80 | 367 | 473 | 0.77 | 472 | 676 | 0.71 | 333 | 484 | 0.67 |
| 1977 | 154 | 234 | 0.66 | 361 | 493 | 0.74 | 471 | 683 | 0.69 | 302 | 456 | 0.78 |
| 1978 | 177 | 211 | 0.82 | 390 | 488 | 0.80 | 452 | 657 | 0.71 | 279 | 437 | 0.72 |
| 1979 | 203 | 203 | 0.97 | 469 | 506 | 0.93 | 508 | 599 | 0.90 | 345 | 482 | 0.67 |
| 1980 | 267 | 215 | 1.22 | 600 | 508 | 1.21 | 763 | 745 | 1.04 | 492 | 553 | 0.86 |
| 1981 | 276 | 218 | 1.28 | 675 | 613 | 1.13 | 695 | 751 | 0.92 | 397 | 520 | 0.75 |
| 1982 | 252 | 238 | 1.05 | 674 | 692 | 0.97 | 543 | 710 | 0.79 | 362 | 509 | 0.69 |
| 1983 | 144 | 152 | 0.95 | 757 | 906 | 0.84 | 421 | 609 | 0.71 | 347 | 538 | 0.64 |
| 1984 | 150 | 156 | 0.95 | 683 | 793 | 0.86 | 370 | 538 | 0.69 | 301 | 463 | 0.63 |
| 1985 | 134 | 141 | 0.91 | 638 | 789 | 0.80 | 407 | 558 | 0.73 | 314 | 481 | 0.65 |
| 1986 | 104 | 109 | 0.94 | 656 | 788 | 0.83 | 368 | 509 | 0.72 | 299 | 468 | 0.64 |
| 1987 | 144 | 143 | 1.01 | 693 | 810 | 0.86 | 459 | 626 | 0.73 | 327 | 518 | 0.63 |
| 1988 | 125 | 122 | 1.02 | 578 | 712 | 0.81 | 361 | 516 | 0.70 | 307 | 504 | 0.61 |
| 1989 | 127 | 124 | 1.02 | 655 | 729 | 0.90 | 396 | 487 | 0.81 | 298 | 408 | 0.73 |
| 1990 | 144 | 139 | 1.02 | 695 | 803 | 0.87 | 405 | 515 | 0.79 | 305 | 440 | 0.70 |
| 1991 | 198 | 176 | 1.11 | 754 | 811 | 0.93 | 527 | 562 | 0.95 | 426 | 475 | 0.90 |
| 1992 | 148 | 120 | 1.21 | 762 | 730 | 1.04 | 429 | 463 | 0.93 | 391 | 431 | 0.91 |
| 1993 | 130 | 112 | 1.13 | 747 | 683 | 1.09 | 403 | 423 | 0.95 | 360 | 404 | 0.89 |
| 1994 | 85 | 72 | 1.14 | 705 | 584 | 1.19 | 532 | 443 | 1.19 | 372 | 395 | 0.94 |
| 1995 | 95 | 74 | 1.25 | 636 | 610 | 1.04 | 540 | 483 | 1.11 | 388 | 410 | 0.95 |
| 1996 | 92 | 78 | 1.17 | 632 | 656 | 0.95 | 500 | 539 | 0.92 | 400 | 471 | 0.85 |
| 1997 | 71 | 59 | 1.20 | 658 | 683 | 0.96 | 524 | 553 | 0.93 | 407 | 449 | 0.91 |
| 1998 | 60 | 45 | 1.28 | 647 | 574 | 1.12 | 550 | 473 | 1.15 | 435 | 430 | 1.01 |
| 1999 | 47 | 27 | 1.70 | 673 | 437 | 1.52 | 552 | 359 | 1.52 | 427 | 330 | 1.30 |
| 2000 | 76 | 42 | 1.79 | 621 | 342 | 1.79 | 568 | 333 | 1.68 | 435 | 312 | 1.41 |
| 2001 | 42 | 21 | 1.97 | 646 | 300 | 2.14 | 570 | 288 | 1.95 | 419 | 284 | 1.49 |
| 2002 | 30 | 15 | 2.01 | 661 | 276 | 2.38 | 570 | 256 | 2.25 | 453 | 287 | 1.58 |
| 2003 | 33 | 16 | 1.95 | 713 | 349 | 2.04 | 640 | 322 | 1.98 | 491 | 346 | 1.42 |
| 2004 | 44 | 18 | 2.42 | 689 | 326 | 2.11 | 615 | 322 | 1.91 | 520 | 371 | 1.40 |
| 2005 | 43 | 20 | 2.14 | 713 | 371 | 1.92 | 631 | 376 | 1.68 | 474 | 401 | 1.18 |
| 2006 | 57 | 32 | 1.77 | 765 | 457 | 1.67 | 606 | 443 | 1.37 | 434 | 399 | 1.09 |
| 2007 | 84 | 53 | 1.57 | 795 | 608 | 1.31 | 538 | 551 | 0.98 | 413 | 439 | 0.94 |
| 2008 | 89 | 80 | 1.11 | 580 | 699 | 0.83 | 375 | 617 | 0.61 | 319 | 489 | 0.65 |
| 2009 | 95 | 115 | 0.82 | 481 | 690 | 0.70 | 360 | 655 | 0.55 | 301 | 580 | 0.52 |
| 2010 | 45 | 51 | 0.88 | 456 | 511 | 0.89 | 437 | 451 | 0.97 | 300 | 302 | 1.00 |
| 2011 | 33 | 43 | 0.78 | 445 | 483 | 0.92 | 411 | 432 | 0.95 | 347 | 321 | 1.08 |
| 2012 | 23 | 38 | 0.62 | 419 | 489 | 0.86 | 399 | 497 | 0.80 | 372 | 379 | 0.98 |
| 2013 | 42 | 46 | 0.92 | 399 | 451 | 0.88 | 411 | 416 | 0.99 | 389 | 336 | 1.16 |
| 2014 | 25 | 25 | 0.88 | 412 | 412 | 1.00 | 384 | 398 | 0.97 | 421 | 367 | 1.15 |
| 2015 | 14 | 12 | 1.20 | 416 | 382 | 1.09 | 386 | 389 | 0.99 | 421 | 429 | 0.98 |
| 2016 | 19 | 14 | 1.42 | 409 | 380 | 1.08 | 393 | 411 | 0.96 | 411 | 486 | 0.85 |
| 2017 | 17 | 17 | 1.45 | 408 | 300 | 1.36 | 415 | 329 | 1.26 | 397 | 374 | 1.06 |
| 2018 | 9 | 6 | 1.47 | 447 | 249 | 1.80 | 405 | 265 | 1.52 | 383 | 319 | 1.20 |
| 2019 | 12 | 8 | 1.60 | 451 | 248 | 1.82 | 371 | 231 | 1.61 | 359 | 267 | 1.35 |
| 2020 | 35 | 23 | 1.53 | 472 | 275 | 1.70 | 411 | 258 | 1.60 | 331 | 205 | 1.61 |

## Appendix 2. Computing confidence intervals for the FIMS indices of abundance

Systematic sampling generally produces more precise estimates than random sampling of natural populations that tend to cluster spatially (Elsayir 2014). However, a universal analytic estimator of confidence interval for systematic samples has not yet been found. In a recent paper (McGarvey et al. 2016), 13 methods for estimating the confidence interval (via the variance of the abundance index estimate) for a systematic sample mean were evaluated. The best performer was $v_{8}$ which was adopted here to compute FIMS confidence intervals.

Sampling error variances of the systematic FIMS legal and undersized catch rate were separately estimated for September and January surveys in each year using the $v_{8}$ formula:

$$
v_{8}= \begin{cases}(1-f)\left(s^{2} / n\right)\left[1+2 / \ln (\hat{\rho})+2 /\left(\hat{\rho}^{-1}-1\right)\right], & \text { if } \hat{\rho}>0 \\ (1-f)\left(s^{2} / n\right), & \text { if } \hat{\rho} \leq 0\end{cases}
$$

where

$$
\hat{\rho}=\frac{1}{s^{2} \cdot(n-1)} \sum_{i=1}^{n-1}\left(x_{i+1}-\bar{x}\right) \cdot\left(x_{i}-\bar{x}\right),
$$

and where $n$ is the number the FIMS pots sampled in each September or January survey, $\left\{x_{i} ; i=1, n\right\}$ denotes the pot-specific measurements of lobster abundance of either legal (nr/potlift) or undersized (nr/potlift) catch rate, $\bar{x}$ is the mean of the index across all potlifts, and $\hat{\rho}$ used in the $v_{8}$ formula is the computed estimate of the correlation between neighbouring potlifts along each FIMS transect. The unbiased sample variance $s^{2}$ is computed in the usual fashion as $s^{2}=\frac{1}{(n-1)} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2}$.

To combine September and January surveys, the reported yearly index of FIMS abundance for legal and undersized lobsters $\left(F I M S_{y}\right)$ is the (unweighted) mean of the computed indices:

$$
F I M S_{y}=\frac{\bar{x}_{\text {spring }}+\bar{x}_{\text {summer }}}{2}
$$

To compute an overall confidence interval for this FIMS yearly index (either legal and undersize), we computed the combined yearly variance using the standard variance formula for a sum and multiplication by a constant:
$\operatorname{Var}\left(\right.$ FIMS $\left._{y}\right)=\operatorname{Var}\left(\frac{\bar{x}_{\text {spring }, y}+\bar{x}_{\text {summer }, y}}{2}\right)=\frac{1}{2^{2}}\left[\operatorname{Var}\left(\bar{x}_{\text {spring }, y}\right)+\operatorname{Var}\left(\bar{x}_{\text {summer }, y}\right)+2 \operatorname{cov}\left(\bar{x}_{\text {spring }, y}, \bar{x}_{\text {summer }, y}\right)\right]$

The $\operatorname{cov}\left(\bar{x}_{\text {spring }, y}, \bar{x}_{\text {summer }, y}\right)$ was computed using the standard covariance formula over all ( $n_{\text {pais }, y}$ ) paired pot locations that were sampled in both spring and summer FIMS surveys of each year, $y$ :

$$
\operatorname{cov}\left(\bar{x}_{\text {spring }, y}, \bar{x}_{\text {summer }, y}\right)=\frac{1}{\left(n_{\text {pairs }, y}\right)^{2}} \sum_{i=1}^{n_{\text {paiss, },}}\left(x_{\text {spring }, y, i}-\bar{x}_{\text {spring }, y}\right) \cdot\left(x_{\text {summer }, y, i}-\bar{x}_{\text {summer }, y}\right) .
$$

The $95 \%$ confidence intervals were computed for each year $y$ as

$$
C I_{95 \%, y}=1.96 \cdot \sqrt{\operatorname{Var}\left(\text { FIMS }_{y}\right)} .
$$

## Appendix 3. Specifications of the qR model including equations, assumptions and model parameters

## Overview

The qR fishery stock assessment model operates on a yearly time-step. It is an age-based model, with a maximum age of $20+$. As data input, it fits to yearly totals for commercial lobster catch in both weight and numbers landed, and conditions on yearly fishing effort. A prior value for instantaneous natural mortality rate is assumed. A vector for mean weight-at-age was estimated from yearly growth increments inferred from tag-recovery data and an assumed length for age-1 lobsters (length of legal recruits).

## Data and fixed parameter inputs

Annual lobster catch in the South Australian lobster fisheries is reported in logbooks by weight $\left(C_{t}^{W}\right)$ and by numbers $\left(C_{t}^{N}\right)$. Effort $\left(E_{t}\right)$ is reported as yearly pot lifts. The model year ( $t=$ 1983, ..., 1983+ $n_{t}-1$ ) runs from 1 June ending 31 May and $n_{t}=$ the number of fishing seasons modelled from 1983 to the most recent year. Age is subscripted by $a$, where $a=1$ refers to lobsters reaching legal minimum length during or in the winter before a given fishing season, and the plus-group age $a=20+$ refers to the highest age group including all lobsters of age 20 years and older. The mean weights-at-age $\left\{w_{a} ; a=1,20+\right\}$ of harvested lobsters (McGarvey et al. 1999) are inputs. An instantaneous natural mortality rate of $M=0.1 \mathrm{yr}^{-1}$ is widely assumed for this species (e.g. Annala and Breen 1989) and genus (Johnston and Bergh 1993).

## The population dynamics model

The qR model is effort-conditioned. A Baranov mortality submodel is assumed, where population number declines exponentially due to mortality within each yearly time step. Recruitment of lobsters to the legal stock in each year is a freely estimated parameter. Catchability is estimated separately for two time periods, before and after the imposition of quota management in 1993.

Model variables are listed in Table 5-3. The array of lobster numbers by age and year, $N_{a, t}$, varies over yearly time due to incoming recruitment, $N_{1, t}=R_{t}$, occurring at the start of each year $t$ and due to outgoing mortality through each year. Natural and fishing mortality were assumed to be independent of age. Growth is expressed in the vector of mean weights at age.

Yearly cohort losses due to natural mortality and harvesting for ages 1-19 years old are written;

$$
\begin{equation*}
N_{a+1, t+1}=N_{a, t} \cdot \exp \left(-Z_{t}\right) \tag{1a}
\end{equation*}
$$

where total instantaneous mortality rate $Z_{t}=F_{t}+M$. For the age 20+ 'plus group', the survival equation is written:

$$
\begin{equation*}
N_{20+, t+1}=\left[N_{19, t}+N_{20+, t}\right] \cdot \exp \left(-Z_{t}\right) . \tag{1b}
\end{equation*}
$$

Deaths due to harvesting were summed over age to yield predicted catches by number ( $\hat{C}_{t}^{N}$ ) and weight ( $\hat{C}_{t}^{W}$ ) for fitting to data in each year of the logbook time series:

$$
\begin{gather*}
\hat{C}_{t}^{N}=\frac{F_{t}}{Z_{t}} \cdot\left\{1-\exp \left(-Z_{t}\right)\right\} \cdot \sum_{a=1}^{20+} N_{a, t}  \tag{2a}\\
\hat{C}_{t}^{W}=\frac{F_{t}}{Z_{t}} \cdot\left\{1-\exp \left(-Z_{t}\right)\right\} \cdot \sum_{a=1}^{20+} w_{a} N_{a, t} \tag{2b}
\end{gather*}
$$

Fishing mortality is assumed to vary in proportion to reported yearly effort, $E_{t}$, related by a catchability coefficient that is different for years before and after quota:

$$
F_{t}=\left\{\begin{array}{ll}
q \cdot E_{t}, & \text { for years prior to quota management }  \tag{3}\\
q^{\text {Quota }} \cdot E_{t}, & \text { for years under quota management }
\end{array} .\right.
$$

The initial population age vector ( $N_{a, 1983}$ ) is estimated assuming a stationary age structure using the first-year estimated recruitment $R_{1983}$ and a freely estimated $F_{0}$ :

$$
\left\{\begin{array}{l}
N_{1,1983}=R_{1983} \\
N_{2,1983}=R_{1983} \exp \left[-\left(M+F_{0}\right)\right] \\
N_{a+1,1983}=N_{a, 1983} \exp \left[-\left(M+F_{0}\right)\right], \quad a=2,19 \\
N_{20+, 1983}=N_{19,1983} \exp \left[-\left(M+F_{0}\right)\right] /\left\{1-\exp \left[-\left(M+F_{0}\right)\right]\right\}
\end{array} .\right.
$$

## Likelihood function

The negative log likelihood is written:

$$
\begin{equation*}
-\log L=n_{t} \log \sigma_{N}+\frac{1}{2 \cdot \sigma_{N}^{2}} \sum_{t=1983}^{1983+n_{t}-1}\left(C_{t}^{N}-\hat{C}_{t}^{N}\right)^{2}+n_{t} \log \sigma_{W}+\frac{1}{2 \cdot \sigma_{W}^{2}} \sum_{t=1983}^{1983+n_{t}-1}\left(C_{t}^{W}-\hat{C}_{t}^{W}\right)^{2} \tag{4}
\end{equation*}
$$

Variances of the two normal likelihood components of Eq. 4 (for catches in number and in weight) were written in terms of a single estimated coefficient-of-variation parameter ( $\sigma_{C}$ ) and the respective data time series means:

$$
\begin{align*}
& \sigma_{N}=\sigma_{C} \cdot \bar{C}^{N}  \tag{5a}\\
& \sigma_{W}=\sigma_{C} \cdot \bar{C}^{W} . \tag{5b}
\end{align*}
$$

Estimates of free parameters, $q, q^{\text {Quota }}, \sigma_{C}, F_{0}$, and of yearly recruit numbers $\left\{R_{t} ; t=1983,1983+n_{t}-1\right\}$, were obtained by minimising the negative log-likelihood using R software environment version 3.6 .2 with function "optimr (option nlminb) of package "optimx". The output indicator of yearly biomass was computed as the sum over all ages of population number by age times mean weight at age. For both LenMod and qR models, biomass is reported as a year-average (rather than start-year) quantity. For qR, where population declines Baranov exponentially through each yearly model time step, year-average biomass is computed by analytically integrating over the negative-exponential survival through each 12month year, giving:

$$
\begin{equation*}
B_{t}=\sum_{a=1}^{20+} w_{a} N_{a, t}\left[1 / Z_{t}\right]\left[1-\exp \left(-Z_{t}\right)\right] . \tag{6}
\end{equation*}
$$

Yearly egg production by female lobsters at the start of each fishing season (in spring) was computed as

$$
\begin{equation*}
\operatorname{Eggs}_{t}=\sum_{a=1}^{20 \pm} m_{a} f_{a} N_{a, t} / 2, \tag{7}
\end{equation*}
$$

where $m_{a}$ and $f_{a}$ are sampled vectors of maturity and fecundity versus age (Prescott et al. 1996), and a sex ratio of one-half was assumed. The unfished level of egg production (UEP)
is computed by setting fishing mortality equal to zero and re-running the qR model dynamics for 2* $n_{t}$ (two times the number of estimated years), taking the final-year value of this unfished equilibrium egg production to be $U E P$. The reference time period for the constant level of recruitment assumed for all years in this zero- $F$ equilibrium $U E P$ run is the mean of historical estimated recruitment over 1990-2011. The yearly percentage of unfished egg production is computed as $\% U E P_{t}=E g g s_{t} / U E P$.

## Catchability

The $q \mathrm{R}$ catchability parameter estimates are $q=2.6 \times 10^{-7}$ potlifts $^{-1}$ and $q^{\text {Quota }}=3.8 \times 10^{-7}$ potlifts ${ }^{-1}$ for pre- and post-TACC management (before 1993 and from 1993 onward), respectively.

## Model adjustments for 2019/20 due to COVID-19 market disruption

A large reduction in access to the Chinese market for South Australian lobster exports occurred when the COVID-19 pandemic response was implemented, starting after 22 January 2020. This induced changes in fishing practices, and lower levels of fishing activity. To prevent this from biasing qR model estimates, only catch rates reported up to 22 January were used as the 2019/20 index of abundance.

To remove the impact of disrupted fishing on the inputs to the qR model, a 'COVID correction' was applied to effectively remove the reported post-22-January 2020 catch rate data from the 2019/20 qR input data set. The qR model consists of three data components for season 2019:
(1) Catch in weight taken as given over the full 8-month fishing season; (2) Effort calculated via COVID-corrected catch rates as described below and; (3) Catch in number which was not corrected after analysis showed that removing the post-22-January data had negligible impact on mean weight. Using a linear model, catch rate in the $S Z$ was corrected by $+1 \%$, meaning the raw reported value was estimated to be about $1 \%$ lower than it would have been had there been no COVID disruption.

The four steps taken to correct catch rates and calculate the yearly effort qR input for the 2019/20 season (using only catch rate information up to 22 January) are detailed as follows. First, two sets of catch and effort data were created, one including the regular non-winter period of data as usually aggregated for qR model fitting (October-May), and a second that includes only catch and effort up to 22 January 2020. Second, a linear statistical model was fitted relating catch rate that includes data only up to 22 January to the data for the regular ('Reg') season catch rate:

## CwPUEReg[iy] ~ aCwPUEReg22Jan + bCwPUEReg22Jan * CwPUE22Jan[iy]. (A3.1)

This linear model was fitted to data from seasons 2003-2018. The 2019/20 season was naturally excluded since the estimated parameters will be used to rescale catch rate and effort for that latest season. The two estimated parameters, the intercept (aCwPUEReg22Jan) and coefficient (bCwPUEReg22Jan) in this linear relationship, quantify the expected regular OctMay season catch rate given catch rate only up to 22 January in a typical season. We chose 2003 onward as the past seasons to be fitted given that average recruitment levels decreased around 2003. The 2010 SZ season was also excluded due to the fact that October was closed. Third, we computed what the regular season catch rate was expected to be for 2019/20 given the catch rate observed up to 22 January:

CwPUECorReg[2019] = aCwPUEReg22Jan + bCwPUEReg22Jan*CwPUE22Jan[2019]. (A3.2)

Finally, we computed the COVID corrected measure of yearly effort for 2019/20 to be taken as input to the qR model:

ECOVIDReg[2019] = CwAlIMonths[2019] / CwPUECorReg[2019]. (A3.3)
where CwAllMonths is the yearly catch weight data over the full 12 months.
The resulting COVID-corrected catch rate is CwPUECorReg[2019] $=1.604$. The raw yearly catch rate in 2019 was 1.586 . The COVID-corrected catch rate was thus $1.17 \%$ higher. Thus we estimate a correction of about $1 \%$ higher for expected full season catch rate inferred when only data up to 22 January 2020 are used compared to the nominal (raw) 2019/20 SZ catch rate. The higher COVID-corrected catch rate value was used to compute yearly effort in qR model estimates for 2019/20.

This correction procedure was also run for mean weight in 2019. The resulting corrected mean weight differed negligibly (by only $0.1 \%$ ) from SZ nominal, and so no correction to mean weight or $q$ R catch in number was needed or applied.

## Incorporating recreational catch

The proportion of catch taken by recreational sector in the SZ is estimated to vary around $3 \%$. Previously, the qR model used only the more reliable commercial catch and effort data. For this assessment and in future, all qR data inputs, yearly totals for catch in weight, catch in number and effort, are scaled upward by the same yearly varying proportion of recreational catch-in-weight used in LenMod.

Table 5-3 Variables of the qR model dynamics and likelihood assessment estimator.

| Model Variable | Description |
| :---: | :---: |
| $a$ | subscript for age, 1 to 20+ (the last age group representing ages 20 years and older) |
| $n_{t}$ | number of fishing seasons modelled |
| $t$ | subscript for yearly fishing season, 1983 to 1983+ $n_{t}-1$ |
| $N_{a, t}$ | number of lobsters of age $a$, at the start of year $t$ |
| $R_{t}$ | estimated number of recruits at start of year $t$ |
| $F_{t}$ | instantaneous fishing mortality rate in year $t$ |
| $q$ | estimated catchability coefficient for pre-quota (pre-1993) years |
| $q^{\text {Quota }}$ | estimated catchability coefficient for years under quota (1993+) |
| $\hat{C}_{t}^{N}$ | model numbers of lobsters caught in year $t$ |
| $\hat{C}_{t}^{W}$ | model weight of catch in year $t$ |
| $N_{t}$ | total population number at start of year $t$ |
| $B_{t}$ | biomass of lobsters averaged across year $t$ |
| Eggs $_{t}$ | eggs produced by female lobsters at start of year $t$ |
| $\sigma_{N}$ | sigma of yearly normal likelihood residuals about model-predicted $\hat{C}_{t}^{N}$ |
| $\sigma_{W}$ | sigma of yearly normal likelihood residuals for data about model-predicted $\hat{C}_{t}^{W}$ |
| $\sigma_{C}$ | estimated coefficient of variation relating $\sigma_{N}$ and $\sigma_{W}$ to data means $\bar{C}^{N}$ and $\bar{C}^{W}$ |
| $F_{0}$ | estimated fishing mortality used to generate the first-year vector of numbers at age |
| $\hat{C}_{t}^{N}$ | model number of lobsters caught in year $t$ |
| $\hat{C}_{t}^{W}$ | model weight of catch in year $t$ |
| $N_{t}$ | total population number at start of year $t$ |
| $B_{t}$ | biomass of lobsters averaged across year $t$ |
| Eggs $_{t}$ | eggs produced by female lobsters at start of year $t$ |
| UEP | unfished egg production, based on average recruitment 1990-2011 |
| \% UEP ${ }_{\text {t }}$ | percentage of unfished egg production in year $t$ |

## Appendix 4. Specifications of the length-structured model (LenMod) including equations, assumptions and model parameters.

## Overview

LenMod is a population dynamics model that operates on a fishing season defined over, for the Southern Zone Rock Lobster Fishery, $T=10$ time-steps (months), starting with an assumed opening of the fishing season in September ( $i=1$ ) to May ( $i=9$ ), with a multi-month June-August ( $i=10$ ) time step covering each winter season. Note that all assessments prior to the 2020/21 season, the first model month was assumed to be October. The duration of the $i^{\text {th }}$ time-step $(i=1, . ., T)$ in units of years is denoted $t_{i}$. Lobster size-classes are in 4 mm bins, the lowest length bin defined as $82.5-86.5 \mathrm{~mm}$ CL, with 29 bins for males and 21 for females. The model population array, $N_{y, i, l}^{s}$, is the number of lobsters by length bin ( $l$ ), sex $(S)$, fishing season ( $y$; hereafter referred to as year), and month ( $i$ ).

## The population dynamics model

## Basic dynamics

The equation that specifies $N_{y, i, l}^{s}$ takes account of natural mortality $M$ (instantaneous yearly rate), fishing mortality, growth, and settlement under the assumption that harvest occurs before growth and settlement:

$$
\begin{equation*}
N_{y, i+1, l}^{s}=\sum_{l^{\prime}} X_{l^{\prime}, l, i}^{s} N_{y, i, l}^{s} e^{-M t_{i}}\left\{1-\tilde{H}_{y, i, l^{\prime}}^{s}\right\}+\Omega_{i}^{s} \Phi_{l}^{s} R_{y} \tag{1}
\end{equation*}
$$

where
$X_{l, l, i, i}^{s}$ is the fraction of the animals of sex $S$ in size-class $l$ that grow into size-class $l$ during time-step $i$;
$\Omega_{i}^{s}$ is the fraction of the settlement that occurs to sex $S$ during time-step $i\left(\sum_{s} \sum_{i} \Omega_{i}^{s}=1\right)$;
$\Phi_{l}^{s}$ is the proportion of the settlement of animals of sex $S$ that occurs to size-class $l$;
$\tilde{H}_{y, i, l}^{s}$ is the exploitation rate on animals of sex $S$ in size-class $l$ at the start of time-step $i$ of year $y$ over all fleets; and
$R_{y}$ is the settlement of animals during year $y$ :

$$
\begin{equation*}
R_{y}=\bar{R} e^{\varepsilon_{y}\left(\sigma_{R, y}\right)^{2} / 2} \tag{2}
\end{equation*}
$$

where $\bar{R}$ is mean settlement, $\varepsilon_{y}$ is the "settlement residual" for year $y, \sigma_{R, y}$ is the standard deviation of the random fluctuations in settlement for year $y$ :

$$
\sigma_{R, y}^{2}= \begin{cases}\tilde{\sigma}_{R}^{2} \tilde{\tau}^{\left(y_{\text {satat }}-y\right)} & \text { if } y \leq y_{\text {start }}  \tag{3}\\ \tilde{\sigma}_{R}^{2} & \text { otherwise }\end{cases}
$$

$\tilde{\sigma}_{R}$ is the extent of variation in settlement for years after $y_{\text {statt }}$, and $\tilde{\tau}$ determines the extent to which $\sigma_{R, y}$ changes with time ( $\tilde{\tau}<1$ means that the settlement will be closer to the mean settlement for the years before $\left.y_{\text {statr }}\right)$.
$B_{y}^{\text {AvgTotLeg }}$ is the reported year-average legal-sized biomass during year $y$, averaging across $T$ months, using start-month population numbers (after half-month natural survival), where $W_{l}^{s}$ is the weight of a lobster of size $l$ and sex $S$ :

$$
\begin{equation*}
B_{y}^{\text {AvgTotLeg }}=\frac{1}{T} \sum_{i=1}^{T} \sum_{s} \sum_{l>=L M L} W_{l}^{s} e^{-M t_{i} / 2} N_{y, i, l}^{s} . \tag{4}
\end{equation*}
$$

Reported yearly exploitation rate is defined as the sum of commercial landed catch in weight data across the $T$ months divided by the year-average legal-sized biomass.

Egg production is given by the following equation for the case in which spawning is assumed to occur at the start of time-step $i_{m}$ of year $y$ :

$$
\begin{equation*}
E g g s_{y}=\sum_{l} m_{l} f_{l} N_{y, i_{s}, l}^{\mathrm{f}} \tag{5}
\end{equation*}
$$

where $m_{l}$ and $f_{l}$ are previously estimated vectors of maturity and fecundity versus length for females in size-class $l, i_{s}$ is the time-step in which spawning occurs ( $i_{s}=$ month 2), and $N_{y, i_{m}, l}^{\mathrm{f}}$ is the total number of females. The unfished level of egg production is computed by setting all estimated parameters to their values (except recruitment) from the stock assessment run, setting catches to zero, and re-running LenMod for 40 years, sufficient to achieve equilibrium.

Recruitment for this zero-catch run is set to the average over the years 1990-2011. The \% of unfished egg production in each year is computed as the ratio of $E g s_{y}$ divided by the final zero-catch equilibrium level of egg production.

## Catches

$C_{y, i}^{f}$ which is the landed catch in weight data by fleet $f$ during time-step $i$ of year $y$. In addition to landed catch, commercial data includes information on spawning lobsters and those brought up dead in the pots, while five surveys (1998, 2001, 2004, 2007, and 2013) are used as the basis to estimate catches for the recreational fleets. $C_{y, i}^{f}$ is used in defining the fully-selected exploitation rate for fleet $f$ during time-step $i$ of year $y, F_{y, i}^{f}$, is calculated as follows:

$$
\begin{equation*}
F_{y, i}^{f}=\frac{\left(1+d_{y, i}^{f}\right) C_{y, i}^{f}}{\sum_{l} \sum_{s} \tilde{S}_{y, i, l}^{s, f}\left(1-\tilde{p}_{i, l}^{s}\right) V_{i}^{s} W_{l}^{s} N_{y, i, l}^{s} e^{-M t_{i} / 2}} \tag{6}
\end{equation*}
$$

where
$d_{y, i}^{f}$ is the ratio of the discarded dead catch to the legal-size catch for fleet $f$ (only for commercials fleets, and is 0 for recreationals);
$V_{i}^{s}$ is the relative sex vulnerability, determined separately for each month $i$, which, if estimated, is either being fixed at a value of 1 for males ( $V_{i}^{\text {males }}=1$ ) and estimated for females, or fixed at a value of 1 for females ( $V_{i}^{\text {females }}=1$ ) and estimated for males; or fixed to 1 for both sexes;
$\tilde{p}_{i, l}^{s}$ is the proportion of mature animals of $\operatorname{sex} S$ in length-class $l$ which are returned live during time-step $i$ because they are spawning ( 0 for males); and
$\tilde{S}_{y, i l}^{s}$ is the vulnerability by length for the gear used on animals of sex $S$ in size-class $l$ during time-step $i$ of year $y$ incorporates the legal minimum size as:

$$
\tilde{S}_{y, i, l}^{s}= \begin{cases}0 & \text { if } L_{l}^{s}+\Delta L_{l}^{s} \leq \mathrm{LML}_{y}  \tag{7}\\ S_{y, i, l}^{s} & \text { if } L_{l}^{s} \geq \mathrm{LML}_{y} \\ S_{y, i, l}^{s}\left(L_{l}^{s}+\Delta L_{l}^{s}-\mathrm{LML}_{y}\right) / \Delta L_{l}^{s} \text { otherwise }\end{cases}
$$

where $\tilde{S}_{y, i, l}^{s, f}=\tilde{S}_{y, i, l}^{s}$ as it is assumed that at any time when recreational fishing takes place the same gear is used as for the commercial fishery. $L_{l}^{s}$ is the lower limit of size-class $l$ for sex
$S, \Delta L_{l}^{s}$ is the width of a size-class $l(4 \mathrm{~mm})$ for sex $S, \mathrm{LML}_{y}$ is the legal minimum size during year $y, S_{y, i, l}^{s}$ is the vulnerability of the gear used on animals of $\operatorname{sex} S$ in size-class $l$. (There were no changes in $\mathrm{LML}_{y}$, which is 98.5 mm carapace length, over the whole time series for the Southern Zone Rock Lobster Fishery.)
$F_{y, i}^{f}$, is used to define $\tilde{H}_{y, i, l^{\prime}}^{s}$ as follows:

$$
\begin{equation*}
\tilde{H}_{y, i, l}^{s}=\sum_{f} \tilde{S}_{y, j, l}^{s}\left(1-\tilde{p}_{i, l}^{s}\right) V_{i}^{s} F_{y, i}^{f} \tag{8}
\end{equation*}
$$

## Catchability

Catchability is estimated separately by month ( $i$ ) and for two time periods, before (1983-1992) and under (1993+) TACC management. Details on the definition of catchability ( $q_{Q, i}^{\mathrm{Conm}}$ ) are given in this Appendix.

Table 5-4 Catchability estimates from LenMod for the SZRLF.

| Month of fishing season | $q_{Q-0, i}^{\text {Comm }}$ <br> $(i)$ | (1983-1992) <br> $q_{Q-1, i}$ <br> Comm <br> $(1993-2020)$ |
| :---: | :---: | :---: |
| September | $\mathrm{N} / \mathrm{A}$ | $\mathrm{N} / \mathrm{A}$ |
| October | $2.7 \times 10^{-7}$ | $3.5 \times 10^{-7}$ |
| November | $2.6 \times 10^{-7}$ | $3.8 \times 10^{-7}$ |
| December | $3.1 \times 10^{-7}$ | $4.5 \times 10^{-7}$ |
| January | $4.3 \times 10^{-7}$ | $6.8 \times 10^{-7}$ |
| February | $4.6 \times 10^{-7}$ | $6.4 \times 10^{-7}$ |
| March | $4.6 \times 10^{-7}$ | $6.4 \times 10^{-7}$ |
| April | $4.6 \times 10^{-7}$ | $6.3 \times 10^{-7}$ |
| May | $\mathrm{N} / \mathrm{A}$ | $2.2 \times 10^{-6}$ |

## Initial conditions

It is impossible to project this model from unexploited equilibrium owing to a lack of historical catch records for the entire period of exploitation. Instead, it is assumed that the population was in equilibrium with respect to the average catch over the first five years for which catches are available in year $y_{\text {start }}-20$. This approach to specifying the initial state of the stock differs from that traditionally adopted for assessments of rock lobster off Tasmania and Victoria (Punt
and Kennedy 1997; Hobday and Punt 2001) in that no attempt is made to estimate an initial exploitation rate. The settlements for years $y_{\text {start }}-20$ to $y_{\text {stat }}-1$ are treated as estimable so that the model is not in equilibrium at the start of year $y_{\text {start }}$.

## The objective function

The objective function summarises the information collected from the fishery and contains contributions from four data sources:
a) Commercial catch rates and independent catch rates,
b) length-sex frequency data from sampling of commercial pot lifts, and
c) commercial catches in number.

## Catch-rate data

The contribution of the catch-rate data for the commercial fishery to the likelihood function is given by:

$$
\begin{equation*}
L_{1 . a}=\prod_{y} \prod_{i} \frac{1}{I_{y, i}^{\text {Comm }} \sqrt{2 \pi} \sigma_{q, Q, i}^{\text {Comb }}} \exp \left(-\frac{\left(\ln I_{y, i}^{\mathrm{Comm}}-\ln \left(q_{Q, i}^{\text {Comm }} B_{y, i}^{e, \mathrm{Comm}}\right)\right)^{2}}{2\left(\sigma_{q, Q, i}^{\text {Comb }}\right)^{2}}\right) \tag{9.a}
\end{equation*}
$$

while the contribution of fishery-independent monitoring survey (FIMS) index data to the likelihood function is given by

$$
\begin{equation*}
L_{1 . b}=\prod_{y} \prod_{i} \frac{1}{K_{y, i}^{\mathrm{FIMS}} \sqrt{2 \pi} \sigma_{q, Q, i}^{\text {Comb }}} \exp \left(-\frac{\left(\ln K_{y, i}^{\mathrm{FIMS}}-\ln \left(\tilde{\tilde{q}}^{\mathrm{FIMS}} q_{Q, i}^{\text {Comm }} B_{y, i}^{e, \mathrm{Comm}}\right)\right)^{2}}{2\left(\sigma_{q, Q, i}^{\text {Comb }}\right)^{2}}\right) \tag{9.b}
\end{equation*}
$$

where
$q_{Q, i}^{\text {Conm }}$ is the commercial catchability coefficient which varies by time-step (month) $i$ and for each of two periods of years namely before (1983-1992) and after (1993+) inception of TACC (differentiated by index $\{Q=0,1\}, 0$ for years prior to quota, and 1 for years under quota);
$I_{y, i}^{\text {Comm }}$ is the catch-rate index for the commercial fleet for year $y$ and time-step $i$;
$\sigma_{q, Q, i}^{\text {Comm }}$ is the standard deviation of the observation error for the commercial fleet for time-step $i$ and for each of two periods of years indexed by $Q$ for before and after inception of TACC;
$\tilde{\tilde{q}}^{\text {HIS }}$ is the FIMS catchability coefficient; and
$K_{y, i}^{\mathrm{FIMS}}$ is the FIMS catch-rate index for time-step $i$ of year $y$.

FIMS catch rates are available since 2005 for the Southern Zone Rock Lobster Fishery for between two or three months surveyed each year, and are derived from sampling pots spaced evenly across transects which span a larger spatial region than that of the concentrated fishing grounds, where catchability by month is assumed to be the same as that for the commercial fishery. The maximum likelihood estimates for $q_{Q, i}^{\text {Comm }}$ and $\sigma_{q, Q, i}^{\text {Comb }}$ were obtained analytically, while the value for $\tilde{\tilde{q}}^{\text {RIS }}$ was estimated as part of the non-linear search procedure.
$B_{y, i}^{e, C o m m}$ is the exploitable biomass available to the commercial fishery (and recreational fishery) midway into time-step $i$ of year $y$ :

$$
\begin{equation*}
B_{y, l}^{e, \text { Comm }}=\sum_{s} \sum_{l} V_{i}^{s}\left(1-\tilde{p}_{i, l}^{s}\right) \tilde{S}_{y, i, l}^{s} W_{l}^{s} e^{-M_{i} / 2} N_{y, i, l}^{s}\left(1-\tilde{H}_{y, i, l}^{s} / 2\right) \tag{10}
\end{equation*}
$$

## Length-frequency data

Length and sex frequency data are available from a sampling program which has been conducted since 1991. This program involves voluntary reporting on the contents of pot lifts by some commercial fishers. The observed fraction, during time-step $i$ of year $y$ by the commercial fishery, of the catch (in number) of animals of $\operatorname{sex} S$ in size-class $l$ (including undersize) is denoted $\rho_{y, i, l}^{s, \text { Comm }}$. The model-estimate of this quantity, $\hat{\rho}_{y, i, l}^{s, \text { Comm }}$, takes account of the vulnerability of the gear and the numbers in each size-class and sex:

$$
\begin{equation*}
\hat{\rho}_{y, i, l}^{s, C o m m}=\tilde{S}_{y, i, l}^{s} V_{i}^{s}\left(1-\tilde{p}_{i, l}^{s}\right) N_{y, i, l}^{s} / \sum_{s^{\prime}} \sum_{l^{\prime}} \tilde{S}_{y, i, l}^{s^{\prime}} V_{i}^{s^{\prime}}\left(1-\tilde{p}_{i, l}^{s^{\prime}}\right) N_{y, i, l^{\prime}}^{s^{\prime}} \tag{11.a}
\end{equation*}
$$

The observed value of $\rho_{y, i, l}^{s, C o m m}$ is assumed to be multinomially distributed, giving the lengthsex frequency likelihood function (ignoring multiplicative constants):

$$
\begin{equation*}
L_{2}=\prod_{y} \prod_{i} \prod_{l} \prod_{s}\left(\hat{\rho}_{y, i, l}^{s, \text { Comm }}\right)^{n_{y, i, l}^{s, C o m m} \omega} \tag{11.b}
\end{equation*}
$$

where $n_{y, i, l}^{s, C o m m}$ is the observed number of lobsters in the sampling program in time-step $i$ of year $y$ of sex $S$ and size-class $l$, and $\omega$ is a down-weighting constant factor to reduce influence of this data relative to the catch-effort data sets (since catch sampling is not random and selectivity is not stationary). Undersize length-sex frequencies are fit as part of the full length-sex frequency data from the sampling program, with the model catch number predictions being proportional to:

$$
\begin{equation*}
S_{y, i, l}^{s} V_{i}^{s}\left(1-\tilde{p}_{i, l}^{s}\right) N_{y, i, l}^{s} l^{-M t_{i} / 2} \tag{12.a}
\end{equation*}
$$

The length-sex frequencies for spawners are also assumed to be multinomial samples, except the model catch number predictions being proportional to:

$$
\begin{equation*}
S_{y, l, l}^{s} V_{i}^{s} \tilde{p}_{i, l}^{s} N_{y, j, l}^{s} e^{-M_{i} / 2} \tag{12.b}
\end{equation*}
$$

## Catch-in-number

The commercial catches in number, $C_{y, i}^{N}$, are assumed to be lognormally distributed. The contribution of these data to the likelihood function is therefore given by:

$$
\begin{equation*}
L_{3}=\prod_{f} \prod_{y} \prod_{i} \frac{1}{C_{y, i}^{N} \sqrt{2 \pi} \sigma_{N}} \exp \left(-\frac{\left(\ell \mathrm{n} C_{y, i}^{N}-\ell \mathrm{n} \hat{C}_{y, i}^{N, \mathrm{Comm}}\right)^{2}}{2 \sigma_{N}^{2}}\right) \tag{13}
\end{equation*}
$$

where $\hat{C}_{y, i}^{N}=\sum_{s} \sum_{l} V_{i}^{s} \tilde{S}_{y, i, l}^{s}\left(1-\tilde{p}_{i, l}^{s}\right) N_{y, i, l}^{s} e^{-M t_{i} / 2} F_{y, i}^{\mathrm{Comm}}$ and $\sigma_{N}^{\text {Conm }}$ is the standard deviation of the observation error in catch numbers, assumed to apply over all time. The spawner discards are also fitted under the assumption that they are lognormally distributed.

## Parameter estimation

Table 5-5 lists the parameters of the population dynamics model and the objective function, and highlights those parameters assumed to be known exactly and those parameters whose values are estimated by fitting the model to the data. Vulnerability-at-length for specified combinations of months is estimated, separately for each sex, by a logistic function of length. Female spawner fractions are based on auxiliary information.

A constraint is placed on the settlement residuals to stabilise the estimation and prevent confounding with mean recruitment. The following term was included in the objective function:

$$
\begin{equation*}
P=0.5 \sum_{y}\left(\varepsilon_{y}\right)^{2} /\left(\sigma_{R, y}^{2}\right) . \tag{14}
\end{equation*}
$$

Estimates of all parameters were obtained by minimising the negative log-likelihood using ADMB (Fournier et al. 2012).

Table 5-5 Parameters of the length-structured model (LenMod) and their sources for the Southern Zone Rock Lobster Fishery.

| Parameter | Description | Value | Sources |
| :---: | :---: | :---: | :---: |
| $\varepsilon_{y}$ | The settlement residuals for year $y$ | Estimated |  |
| $\ell \mathrm{n}(\bar{R})$ | Mean settlement | Estimated |  |
| $\tilde{\sigma}_{R}$ | The extent of variation in settlement for years after $y_{\text {start }}$ | 0.5 | Assumed |
| $\tilde{\tau}$ | The extent to which $\sigma_{R, y}$ changes with time | 0.8 | Assumed |
| $X_{l, l, i}^{s}$ | Growth transition matrix | Matrices by sex for months 4 and 9. | Estimated using method of McGarvey and Feenstra (2001). |
| M | Natural mortality | $0.1 \mathrm{yr}^{-1}$ | Conventional assumption |
| $V_{i}^{s}$ | Relative vulnerability by sex by time-step | Fixed at 1 for all months and both sexes. |  |
| $S_{y, i, l}^{s}$ | Vulnerability of the gear by sex, size-class, time-step, and year. | Estimated as logistic functions of length per sex, shared across years, but separately for Sep-Dec, Jan-March, April, and May. |  |
| $\tilde{p}_{i, l}^{s}$ | Proportion of mature spawning animals by sex, size-class and time-step |  | Estimated externally |
| $\Omega$ | Fraction of the settlement by time-step and sex | Estimated |  |
| $\Phi_{l}^{s}$ | Proportion of the settlement of animals by sex and sizeclass | $\begin{aligned} & \text { First six length bins: males }=0.35,0.2 \text {, } \\ & 0.15,0.15,0.1,0.05 ; \text { females }=0.45, \\ & 0.25,0.15,0.1,0.05,0 \end{aligned}$ | Assumed |
| $Q_{l}$ | Egg production as a function of size |  | Estimated externally |
| $W_{l}^{s}$ | Mass as a function of size and sex | Power function of length | Estimated externally |


| $i_{\mathrm{m}}$ | The time-step in which spawning occurs | 2 |
| :--- | :--- | :--- |
| $q_{Q, i}^{\text {Comm }}, \tilde{\tilde{q}}^{\text {HMS }}$ | Catchability for the commercial fleet and FIMS by time-step <br> $i$ and for each of two periods of years namely before and <br> after inception of TACC <br> Standard deviation of the observation errors for time-step $i$ <br> and for each of two periods of years namely before and <br> after inception of TACC for the commercial fleet, and FIMS <br> surveys combined after TACC inception. | Estimated |
| $\sigma_{N, Q, i}^{\text {Comm }}$ | Standard deviation of the observation error in commercial <br> catch in numbers | Estimated |
| $\omega$ | Down-weighting factor for length-sex data | 0.0125 |

## Appendix 5. Model fits

## qR model



Figure 5-1 Fit of the qR model to catch in number of lobsters landed for the SZRLF, based on annual logbook catch totals from the fishery.


Figure 5-2 Fit of the qR model to catch in weight for the SZRLF, based on annual logbooks catch totals from the fishery.

LenMod


Figure 5-3 Fit of the LenMod model to monthly catch in number (Cn) for the SZRLF, based on logbook catch totals from the fishery. (Note: October closed to fishing in 2010).


Season
Figure 5-4 Fit of the LenMod model to monthly catch per unit effort (CPUE) for the SZRLF, based on logbook catch totals from the fishery (Note: October closed to fishing in 2010).


Figure 5-5 Fits of LenMod model (black line) proportions by length bin to commercial length frequency data for both males and females taken during the 2020 season in the SZRLF.

